

CH 17

便箋
標題

2008/5/29

* Note on Convergence Concept

We introduce five convergence Modes involving sequences of random variables.

- Def: 1) random sequence is a sequence of

random variables $X_1(s)$, $X_2(s)$, ...
outcome

For a given outcome s_0 , $X_1(s_0)$, $X_2(s_0)$, ... is a deterministic sequence which may or may not converge.

- Converged Everywhere (c)

A random sequence is said to converge everywhere $X(s)$ if, given $\epsilon > 0$, we can find a number n_0 , such that

$$|X_n(s) - X(s)| < \epsilon, \quad \forall s \left(\lim_{n \rightarrow \infty} X_n(s) = X(s) \right)$$

for every $n > n_0$. We use

$$\lim_{n \rightarrow \infty} X_n(s) = X(s) \quad \text{or} \quad X(s)$$

To denote this mode.

- Convergence Almost Everywhere (a.e) $\{s | \lim_{n \rightarrow \infty} X_n(s) = X(s)\}$ exist and if the event $\{s | \lim_{n \rightarrow \infty} X_n(s) = X(s)\}$ has probability 1, then we say that the

segment converges almost everywhere to $X(s)$ (or with probability one). We denote the

$$\text{mode by } \mathbb{P} \left[\lim_{n \rightarrow \infty} X_n(s) = X(s) \right] = 1.$$

• Convergence in the Mean Square (MS)

Sense

The segment $X_1(s), X_2(s), \dots$ converges to $X(s)$

in the MS sense of

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[|X_n(s) - X(s)|^2 \right] = 0.$$

We denote the mode by

$$\text{L.i.m. } X_n(s) = X(s)$$

Where $L.i.m.$ means "limit in the ms sense."

• Convergence in Probability (P)

$$\text{if } \lim_{n \rightarrow \infty} P[\|X_n(s) - X(s)\| > \epsilon] = 0$$

for any $\epsilon > 0$, then the random sequence

$X_1(s), X_2(s), \dots$ converges to $X(s)$ in probability.

• Convergence in Distribution (D)

$$\text{if } \lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x)$$

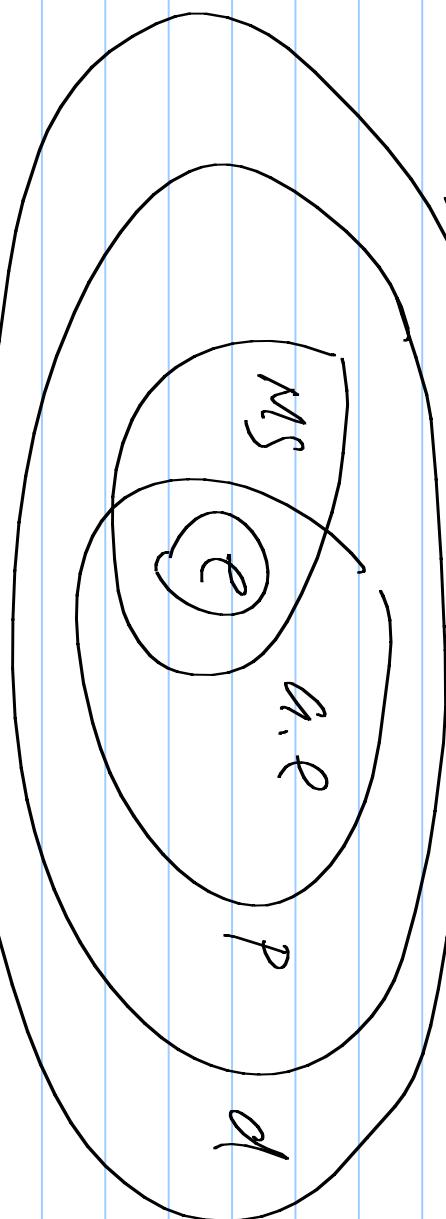
for every point x of continuity of $F_X(x)$,
then the random sequence $X_1(s), X_2(s), \dots$

converges to $X(s)$'s distribution.

"Note: The relationship between various convergence

modes is depicted by

$$\begin{array}{c} \vartriangleleft \\ P \Rightarrow a.e \Rightarrow D_s \Rightarrow q \\ \Downarrow \\ MS \Rightarrow P \Rightarrow d \end{array}$$



Mits' Central limit theorem is defined in
the distribution sense,

② Laws of Large Number are defined in
the probability sense (weak) or in
probability-one sense (strong).

Parameter: mean, moments

Consider an event A . We define the
indicator random variable
 $X_A = \begin{cases} 1 & \text{if event } A \text{ occurs} \\ 0 & \text{otherwise.} \end{cases}$

X_A is a Bernoulli r.v. with success probability
 $P[A]$ and $E[X_A] = P[A]$.

Consider a iid random sequence X_1, X_2, \dots
 X_1, X_2, \dots

The sample mean

$$\hat{M}_n(\hat{x}) = \frac{\sum_{i=1}^n \hat{x}_i}{n} = \overbrace{\hat{P}_n(A)}$$

The segment $M_1(x), M_2(x), \dots$ may converge
in certain sense.

Define R_n to a function of x_1, x_2, \dots, x_n

R_n is used to serve as a point estimate
for certain parameter of the probability
model for \hat{x} .

Consider $\gamma = \text{Var}(x)$.

If $E[x] = 0$, $\text{Var}[x] = E[\bar{x}^2]$. We can

define $y_n = x_n^2$. Then,

$$M_n(y) = \frac{1}{n} \sum_{i=1}^n y_i = \frac{1}{n} \sum_{i=1}^n x_i^2 \rightarrow M_n(x)$$

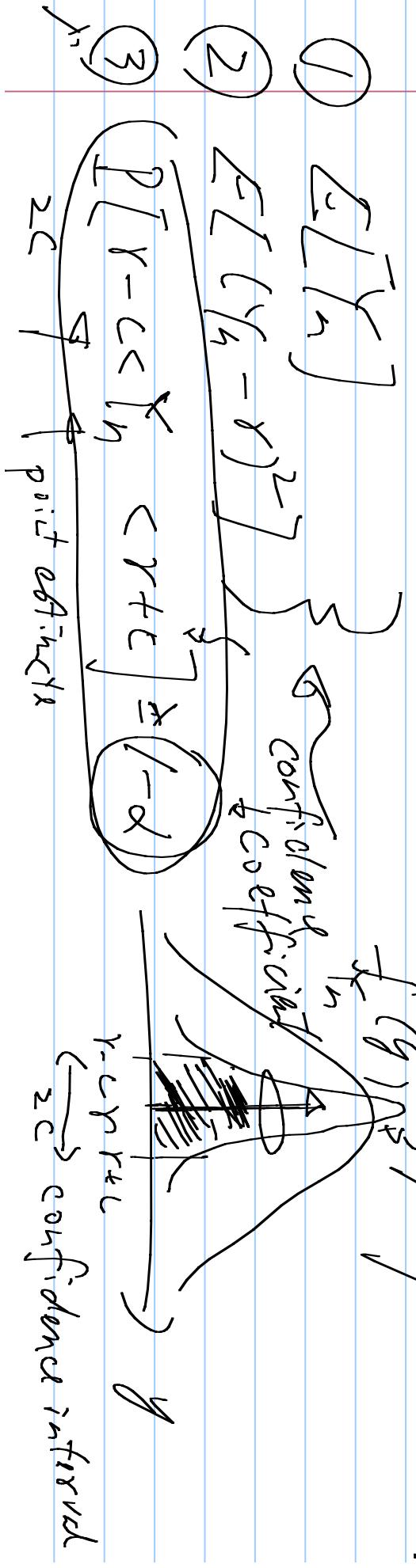
converges in probability to $E[\bar{x}^2] = \text{Var}(\bar{x})$ (according to the weak law of large numbers).

If $E[x] = \mu_x$, then we can form $y_n = (x_n - \mu_x)^2$ and show that $(M_n(y))$ converges to $\text{Var}(\bar{x})$ in probability.

$$\frac{1}{n} \sum_{i=1}^n y_i = \frac{1}{n} \sum_{i=1}^n (x_i - E[x])^2$$

Consider x_1, x_2, \dots, x_n which are iid with known probability density form, however with unknown parameters θ from $y_n = g(x_1, x_2, \dots, x_n)$ to estimate θ estimate of the unknown parameters.

Characterize the estimation performance.



$$(4) \quad P[\underline{Y_n} - c < \bar{Y}_n < \underline{Y_n} + c] \geq 1 - \alpha$$

Confidence Interval = $[\underline{Y_n} - c, \underline{Y_n} + c]$ is random

$$P[\bar{Y}_n \in (\underline{Y_n} - c, \underline{Y_n} + c)] \geq 1 - \alpha$$

The length of confidence interval is $2c$, which is deterministic.

$\exists x$: Consider the sample mean $\bar{Y}_n(x)$ so as to obtain c of $E(\bar{Y}_n)$. From Chebyshev's inequality,

$$P[\left| \bar{Y}_n(x) - E(\bar{Y}_n) \right| \geq c] \leq \frac{V(\bar{Y}_n(x))}{c^2} = \frac{V(\bar{Y}_n)}{c^2}$$

In terms of "interval estimation" formulation

$$\Pr[M_n(x) - c \leq E[\bar{X}] \leq M_n(x) + c] = \frac{\text{Var}[\bar{X}]}{n c^2}$$

where

$\Pr[M_n(x) - c \leq M_n(x) + c]$ no the random confidence interval with fixed length $2c$

(2). $\frac{\text{Var}[\bar{X}]}{n c^2}$ no the confidence coefficient

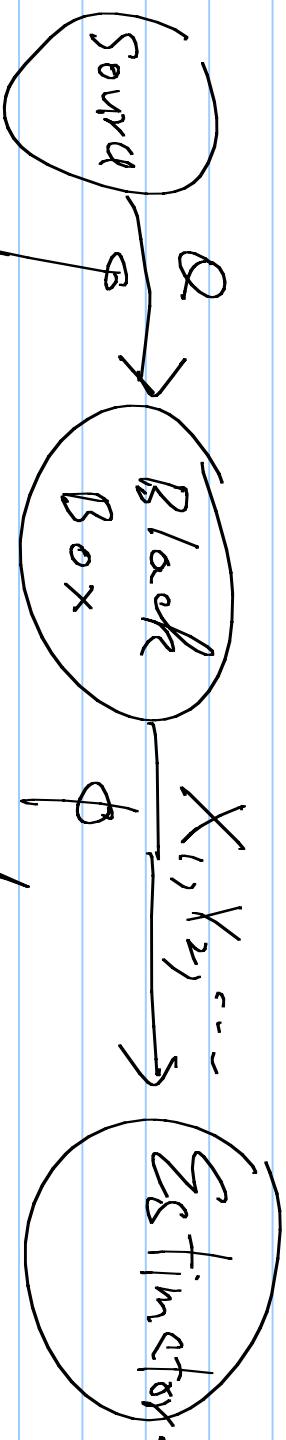
sinus

$$\Pr[E[\bar{X}] \notin \{M_n(x) - c, M_n(x) + c\}] \geq 1 - \frac{\text{Var}[\bar{X}]}{n c^2}$$

$1 - \alpha$

Parameter Estimation

$$\hat{\theta} = f(x_1, x_2, \dots)$$



parameter estimator of θ
with known
probability variables

model identical

$x_i \sim$ Erlang and Gaussian with known θ
and identical variance 1. A good estimate
is the sample mean, i.e., $\hat{\theta} = \frac{1}{n} \sum_{i=1}^n x_i$

$$\frac{dy}{dx} = \left[(x)^y \right]_{15}^{\infty} \quad \left[x \right]_{15}^{\infty} = \left[(\cancel{x})^{\cancel{y}} \right]_{15}^{\infty}$$

or $y(x) \approx 1$ for $x \gg 1$