

CHAPTER TWO

便箋標題

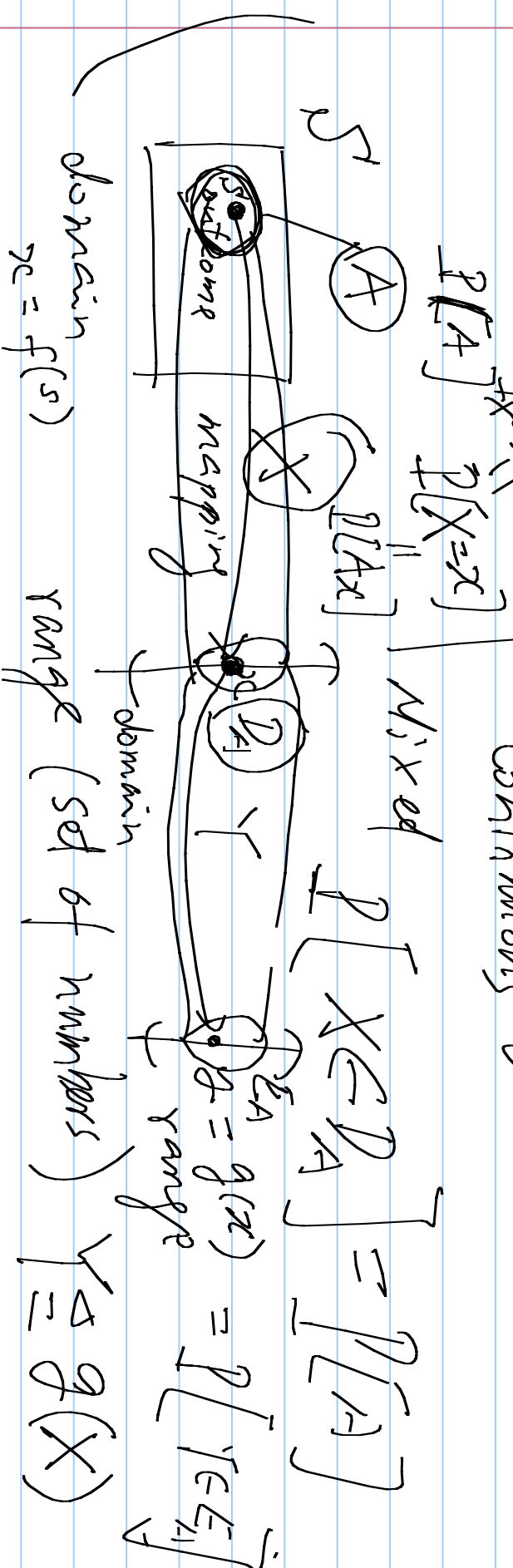
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Experiment: Procedure and Observation

Physical Model \Rightarrow Probability Model

Random Variable

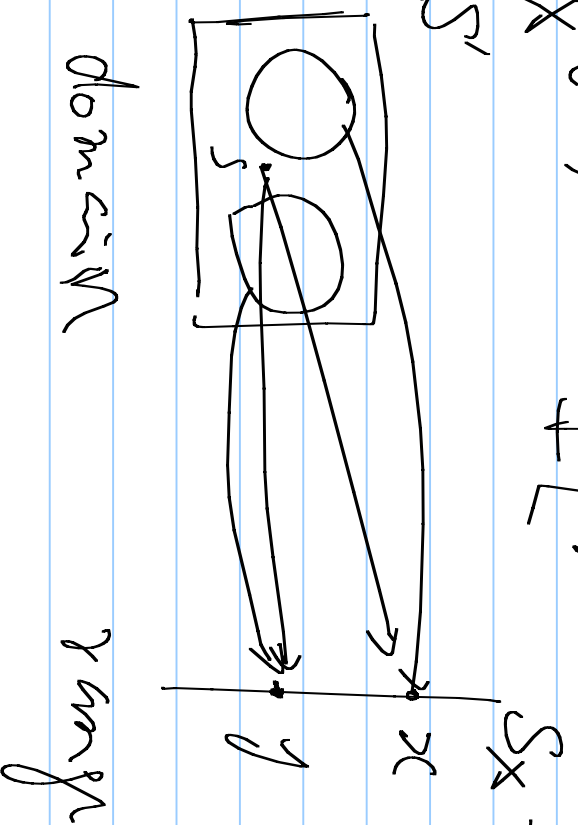
Discrete \checkmark
Continuous \checkmark



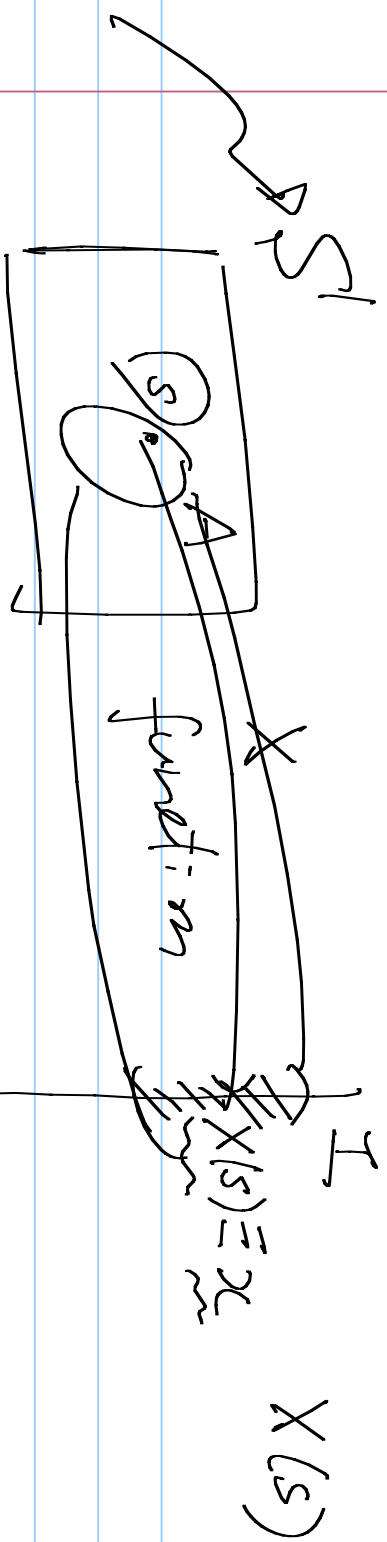
$$\boxed{P[X=x]} \equiv P[A] \text{ where}$$

$$A = \{s \mid X(s) = x\}$$

$$P_X(x) \equiv P[X=x]$$



$$\{X=x\} \cap \{X=y\} = \emptyset \text{ if } x \neq y.$$



P.M.F

Domain

Range

$$P_X(x) \equiv \underbrace{P\left[\underbrace{X(s)=x}_{\text{Range}}\right]}_{\text{Domain}} = P\left[\underbrace{X(s)=x}_{\text{Range}}\right]$$

$$P\{X \in I\} = P\{A\}$$

$$P\{X \in I\} = P[A]$$

A discrete random variable X is said to be completely statistically characterized by its P.M.F.

X

$X \ X \ \dots \ X$

(r)

$(X_{i+1}) - (k-1)$

$r \text{ or } s$

$(k-1)$

r

(X_{i+1})

p^{k-1}

$(1-p)$

p

$(X_{i-(k-1)})$

r

$r \ a$

$k-1$ success

$q \ q \ \dots \ q \ r$

$q \ q \ r \ q \ \dots \ r$

(r)

(≤ -1)
 $(k-1)$
 $p \ p^{k-1} \ (1-p)^{\leq -k}$

Recall: Consider ^{Repeatable} n experiment with identical
and independent subexperiments.
Each subexperiment gives two possible
outcomes: namely success, with
probability p , and failure, with
probability $1-p$.

Bernoulli (p) random variable represents the
number of successes in one single
trial

Geometric (p) random variable represents the
number of trials until the first.

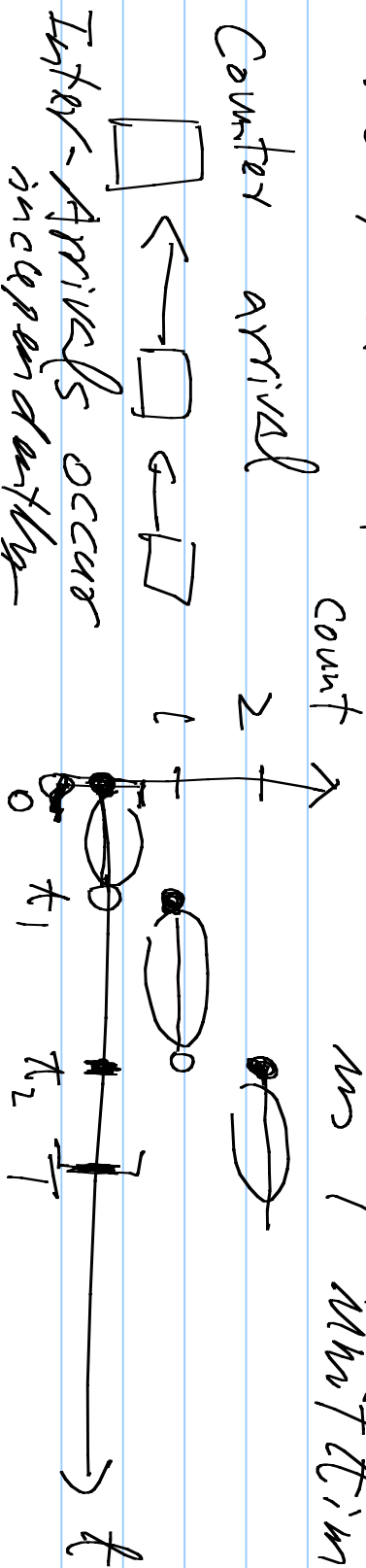
Success (inclusive)

Binomial (n, p) random variable represents

The number of successes in n trials.
Note: Binomial $(1, p) = \text{Bernoulli}(p)$
Pascal (k, p) random variable represents

the number of trials until k successes (inclusive). Note: Pascal $(1, p)$ is equivalent to Geometric (p) .

Poisson (λ) Random Variable = number of arrivals in T unit times



α denotes the average number of arrivals
in T unit times

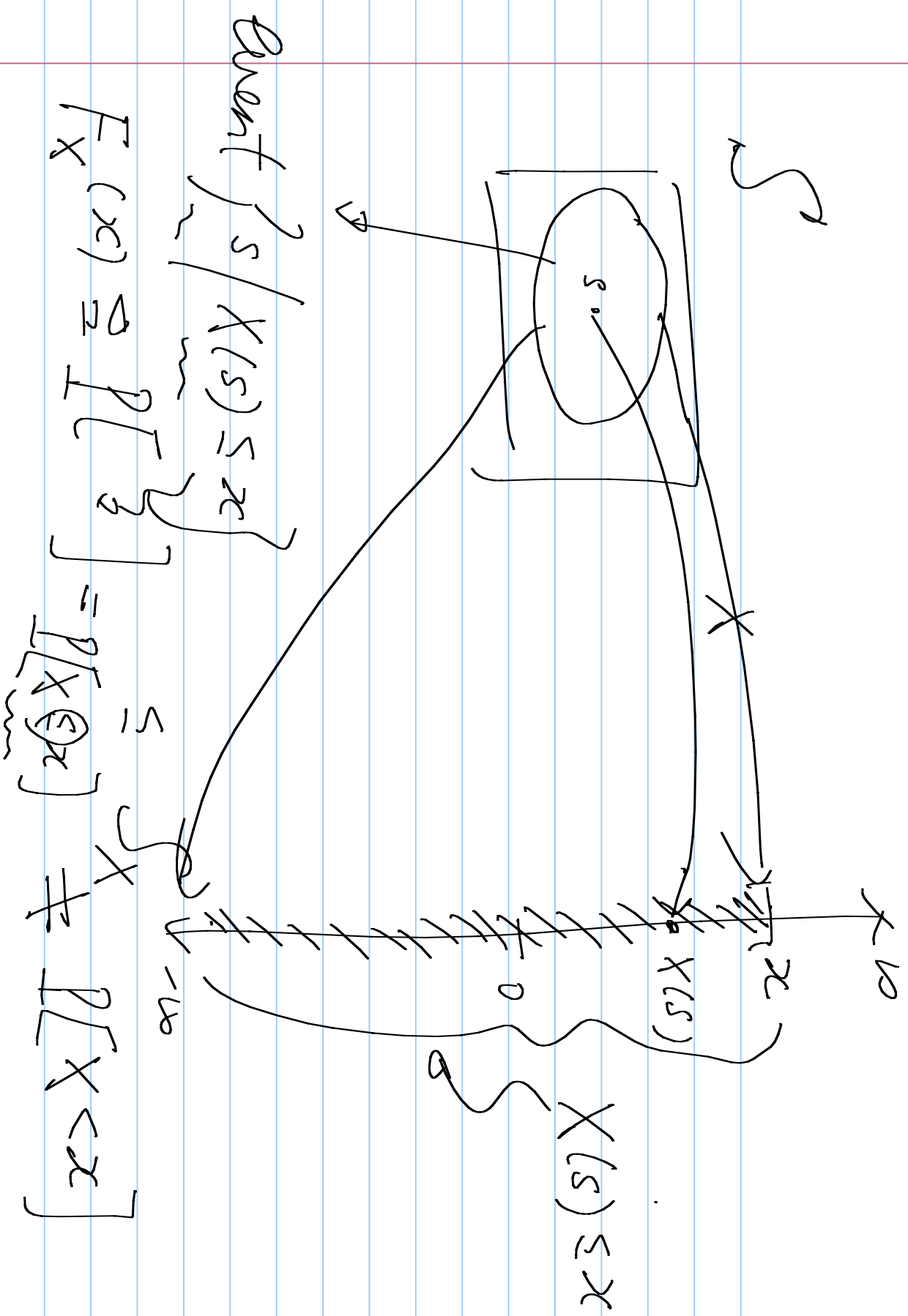
λ denotes the average number of arrivals
in unit time

$$P_k(R) = \frac{\alpha^k e^{-\alpha}}{k!}, k=0, 1, \dots$$

PMF describes a complete probability

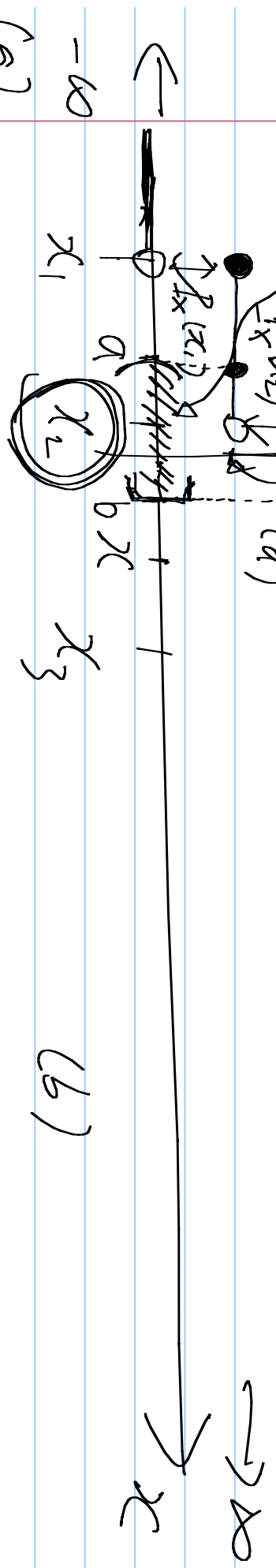
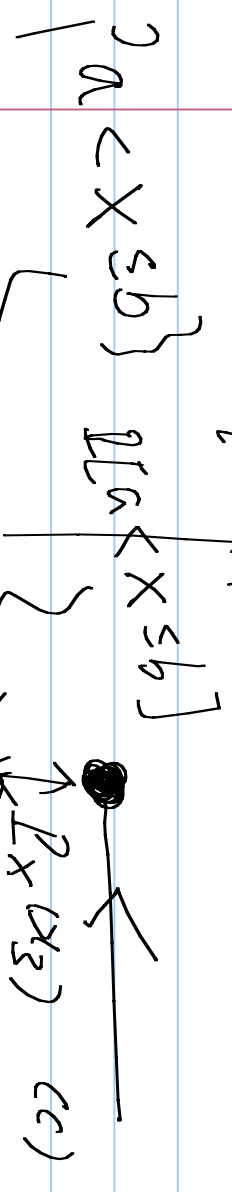
model of a discrete random variable.

(Cumulative) Distribution Function (CDF)
also describes a complete probability model.



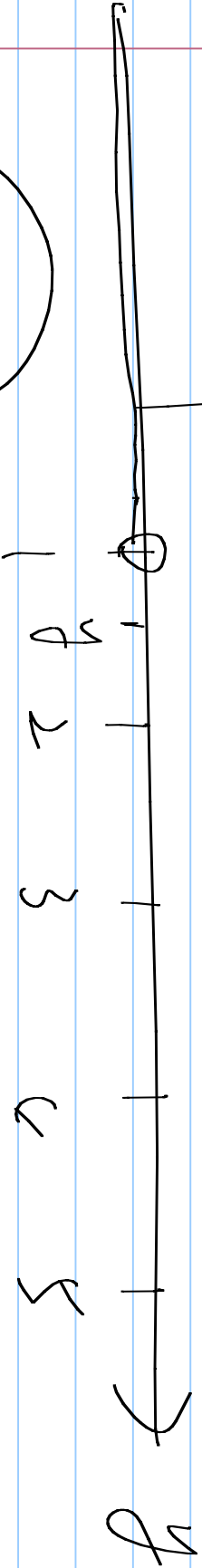
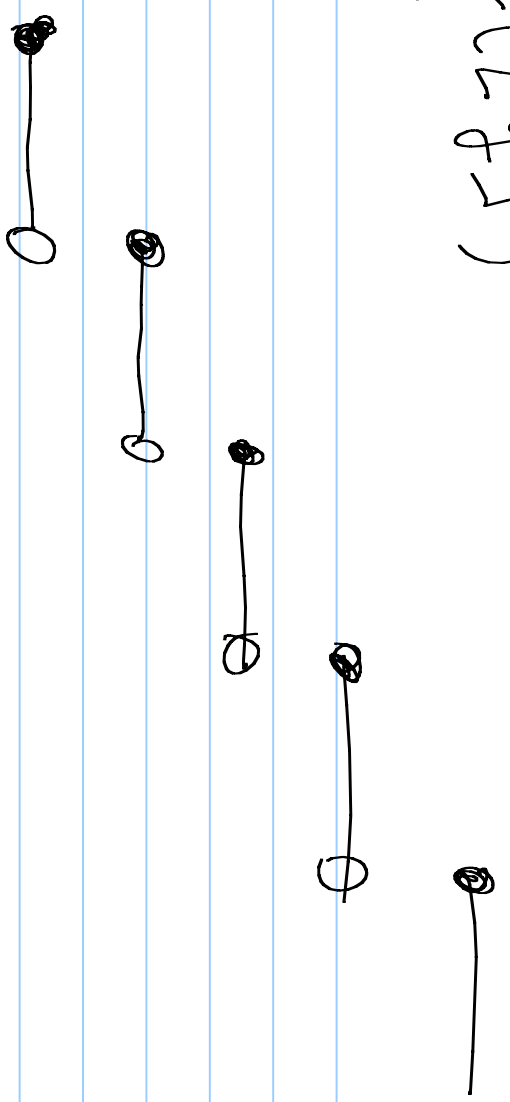
$F_X(x)$ must be nondecreasing. $S_X = \{x_1, x_2, \dots\}$

$F_X(x) \equiv P[X \leq x]$



(c) $F_X(x)$ is continuous from the right, but not necessarily from the left.

$$F_r(y) = F_r(\lfloor y \rfloor)$$



$\lfloor y \rfloor$ is the largest integer that is not larger than y .
 is integer part of y .

[y] \equiv the smallest integer that is not smaller than y

A statistic is a single number derived from (numerical)

• The average value of a set of n experimental outcomes is a statistic of the outcomes.

Three averages of interest:

mean (or expected value, or expectation)

median

mode (unimodal, multimodal)

Consider an experiment that produces a rv X .
Perform n independent trials of this
experiment. Let x_i denote the i th
sample value.

$$M_n \equiv \text{sample average} = \frac{1}{n} \sum_{i=1}^n x_i.$$

$$= \frac{1}{n} \sum_{x \in S_X} x N_x \quad (N_x \equiv \text{the number of sample values } x \text{ in } n \text{ trials})$$

$$M_n = \sum_{x \in S_X} x \frac{N_x}{n}$$

We can interpret $P[X=x]$ as $\lim_{n \rightarrow \infty} \frac{N_x}{n}$
(relative frequency interpretation)

$$\begin{aligned}
 \lim_{n \rightarrow \infty} M_n &= \lim_{n \rightarrow \infty} \sum_{x \in S_X} x \frac{d(x)}{n} \\
 &= \sum_{x \in S_X} x \underbrace{\lim_{n \rightarrow \infty} \frac{1}{n}}_{P[X=x]} \\
 &= \sum_{x \in S_X} x P[X=x] \\
 &= E[X]
 \end{aligned}$$

$$\sum_{x=0}^{\infty} r^x = \frac{1}{1-r} \quad \text{for } 0 < r < 1$$

$$\sum_{x=0}^{\infty} x r^x = \sum_{x=0}^{\infty} r^x = \sum_{x=0}^{\infty} r^x$$

$$\sum_{x=1}^{\infty} x r^x = \sum_{x=0}^{\infty} x r^x = \sum_{x=0}^{\infty} \frac{d}{dr} r^x = \frac{d}{dr} \sum_{x=0}^{\infty} r^x = \frac{d}{dr} \frac{1}{1-r} = \frac{1}{(1-r)^2}$$

$$\frac{1}{1-r} + \frac{1}{(1-r)^2}$$

$$\frac{r}{(1-r)^2} = \frac{r}{(1-r)^2} = \frac{1}{(1-r)^2}$$

$$\frac{1}{(1-r)^2} = \frac{1}{(1-r)^2}$$

$$\lim_{n \rightarrow \infty} P_k(R) = \lim_{n \rightarrow \infty} \frac{n(n-1) \dots (n-k+1)}{n^k}$$

$$\cdot \left(\frac{R}{R} \right)^k \cdot \lim_{n \rightarrow \infty} \left(1 - \frac{\alpha}{n} \right)^{n-k}$$

$$\frac{1}{R^k} \cdot 1 = 1$$

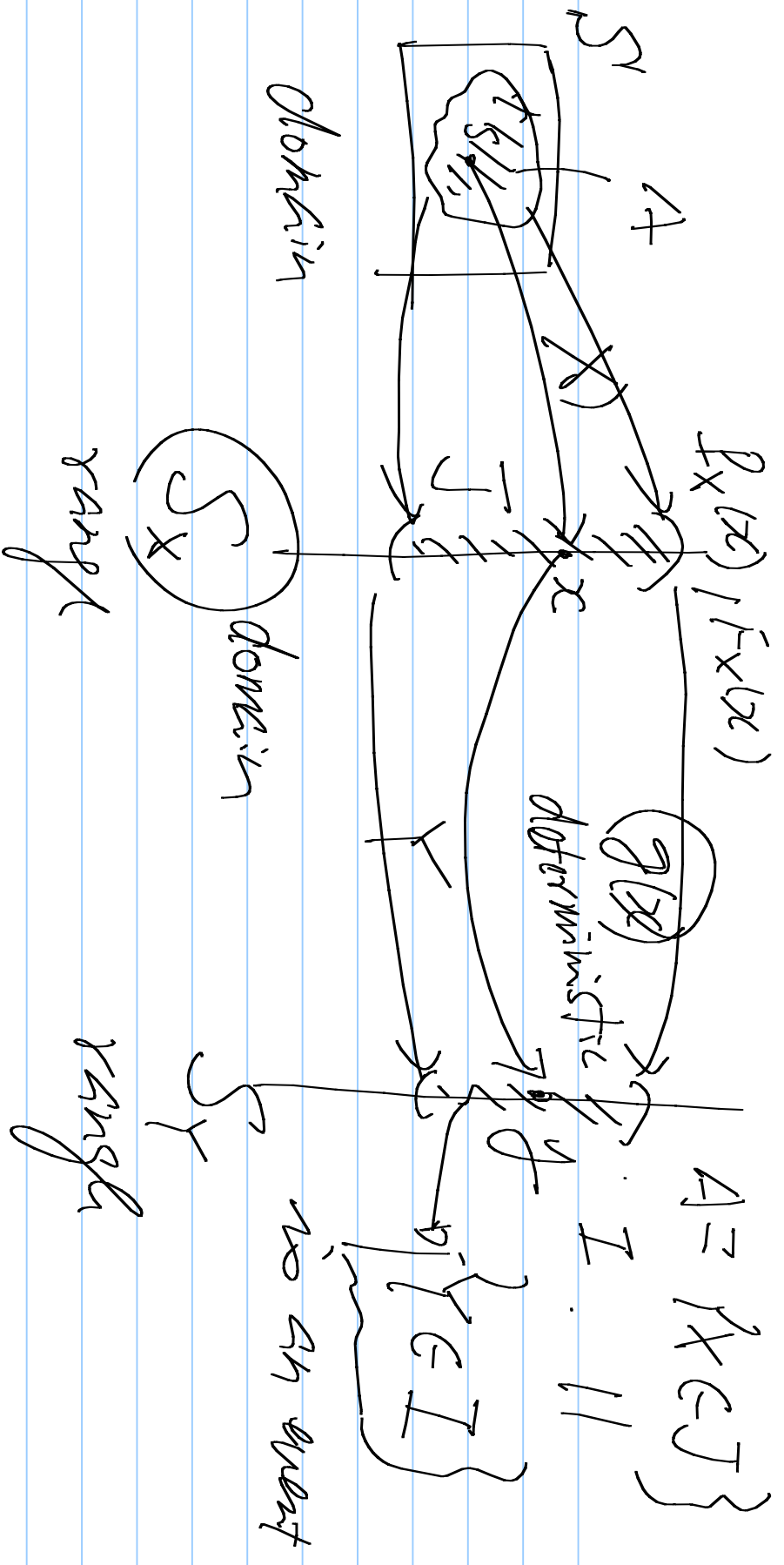
$$\lim_{n \rightarrow \infty} \left(1 - \frac{\alpha}{n} \right)^n = \lim_{n \rightarrow \infty} \left(1 - \frac{\alpha}{n} \right)^k = 1$$

$$\lim_{n \rightarrow \infty} \left(1 - \frac{\alpha}{n} \right)^k = 1$$

$$= \lim_{n \rightarrow \infty} \left(1 - \frac{\alpha}{n} \right)^n = \lim_{n \rightarrow \infty} e^{n \ln \left(1 - \frac{\alpha}{n} \right)}$$

$$\frac{0}{0}$$

$$\begin{aligned}
 L'_{\text{Hopital}} &= \lim_{h \rightarrow 0} \frac{h \ln(1 - \frac{\alpha}{h})}{\frac{\alpha}{h}} = \lim_{h \rightarrow 0} \frac{\ln(1 - \frac{\alpha}{h})}{1/h} \\
 &= \lim_{h \rightarrow 0} \frac{\frac{1 - \frac{\alpha}{h}}{-h^{-2}}}{-1/h} = \lim_{h \rightarrow 0} \frac{-\alpha}{1 - \frac{\alpha}{h}} = -\alpha
 \end{aligned}$$



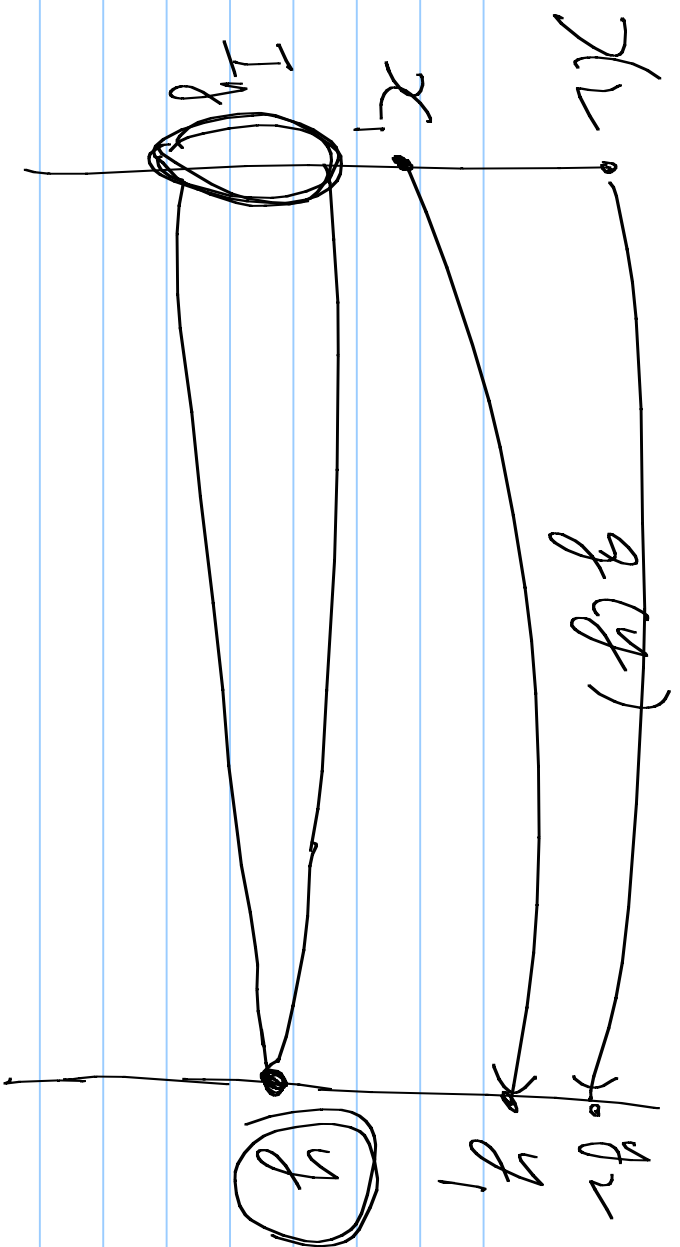
X is completely statistically characterized

if $P_X(x)$ or $F_X(x)$ is known.
 $A = \{s | X(s) \in J\} = \{s | Y(s) \in I\}$

Y is called the derived random

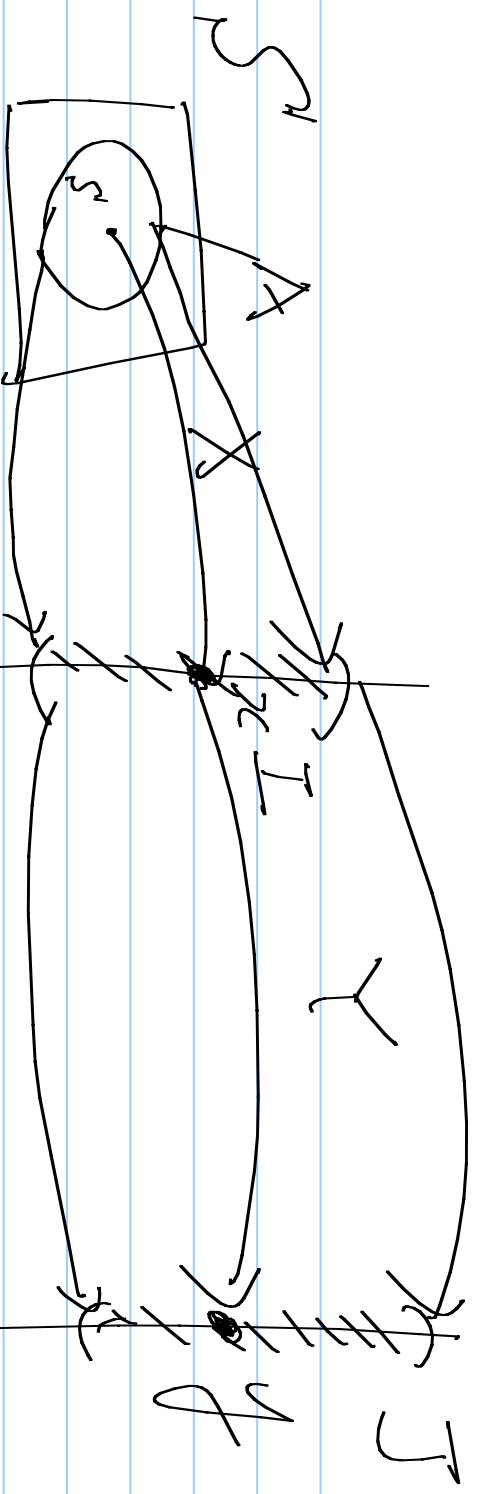
variable.

Denote $Y = g(X)$ ✓



- ① When $g(x_1) \neq g(x_2)$ for $x_1 \neq x_2$ (i.e., $g(x)$ is one-to-one), $P_Y(y) \stackrel{!}{=} P_X(x)$ if $y = g(x)$.
- ② When $y = g(x)$ for all $x \in I_y \stackrel{!}{=} \{x \mid g(x) = y\}$.

$$P_Y(y) = \sum_{x \in I_y} \underbrace{P_X(x)}$$



domain

$$A = \left\{ s \mid X(s) \in I \right\} \quad S_X$$

range

$$= \left\{ X \in I \right\}$$

$$= \left\{ Y(X(s)) \in J \right\} \quad \text{domain}$$

$$= \left\{ Y \in J \right\}$$

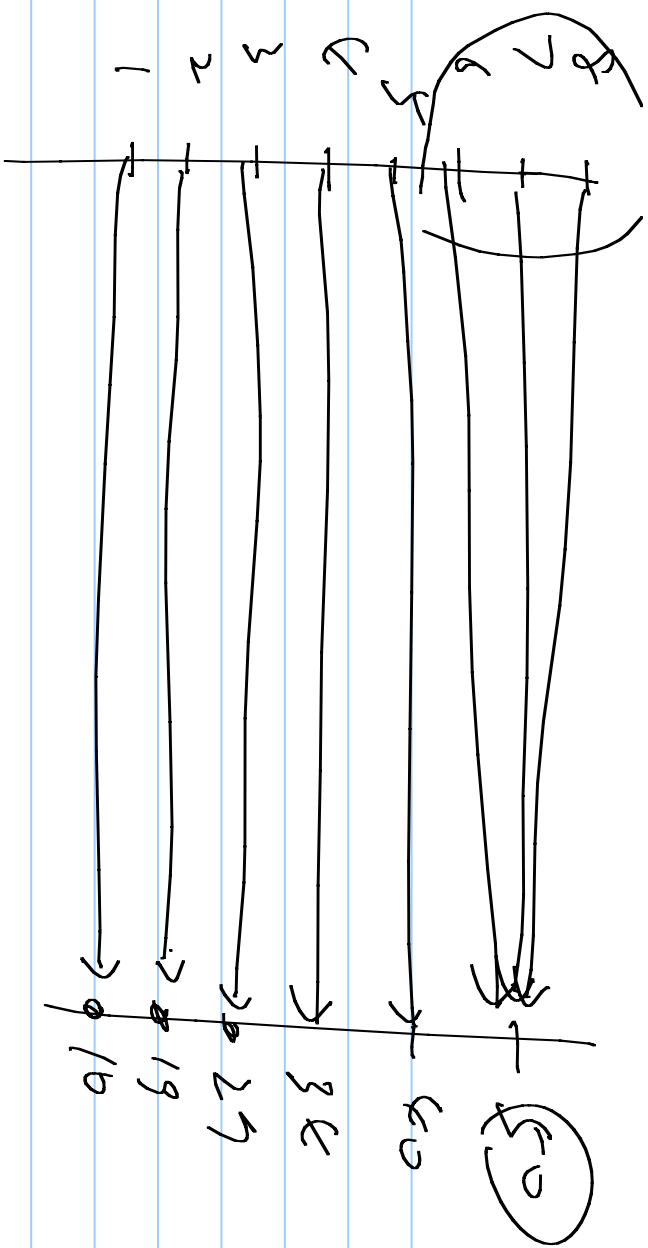
range

$$\left(Y = g(X) \right)$$

When $\{ \underbrace{I = y} \}$, $A = \{ \cancel{S} \mid g(x) = y \} = \{ S \mid \underbrace{g(x(s)) = y} \}$ and

$$\underbrace{P[Y = y]} = P[Y = y] = \underbrace{P[A]} = \sum_{x: g(x) = y} P_X(x)$$

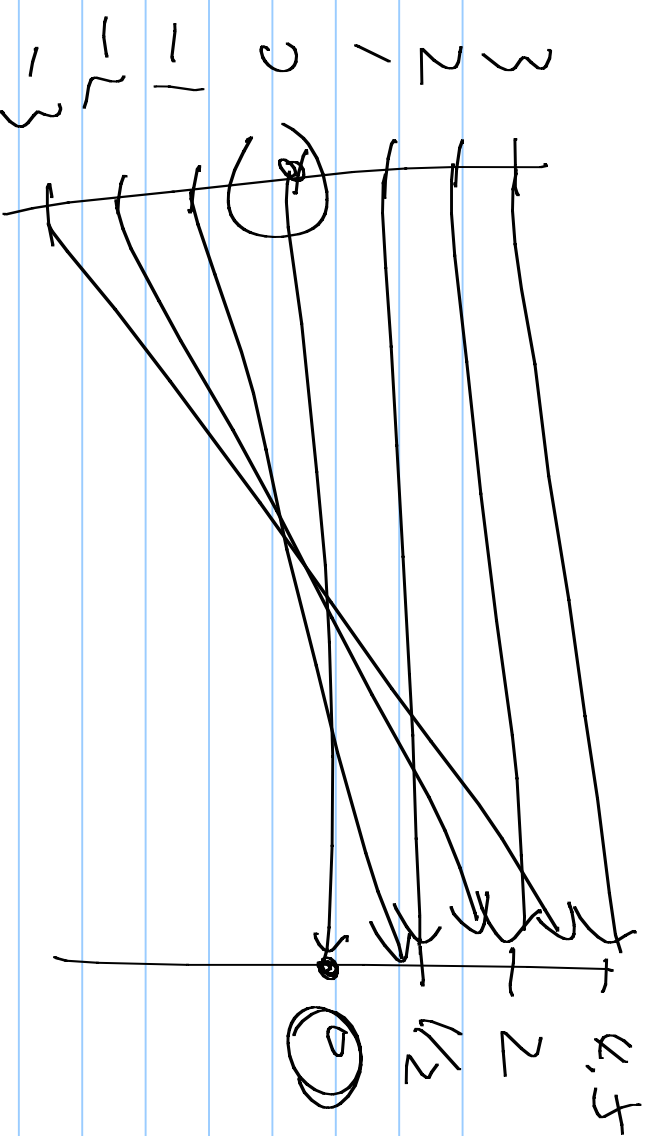
$$P_Y(y)$$



S_x S_y

Many - to - one mapping

Given $P_x(x)$ and $y = g(x)$, this section deals with the approach of finding $P_y(y)$.

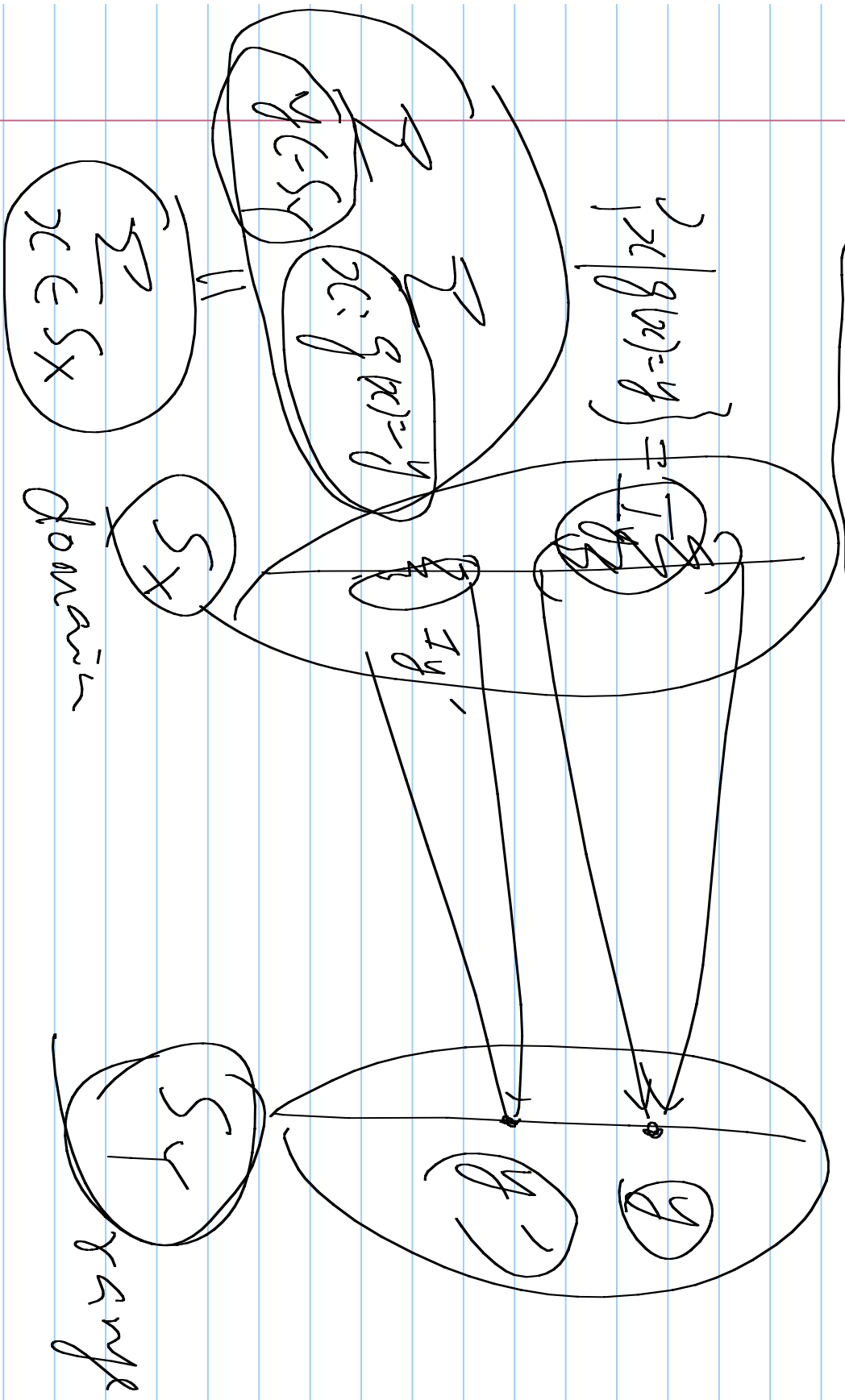


S_X

S_Y

$$\left. \begin{aligned} \{Y=0\} &= \{X=0\}, \{Y=\frac{1}{2}\} = \{X \in]-1, 1[\} \\ \{Y=2\} &= \{X \in]-2, 2[\}, \{Y=4.5\} = \{X \in]-3, 3[\} \end{aligned} \right\}$$

Law of Unconscious Statistician



* An operator $\mathbb{L}[x]$ is called

Linear (affine) iff for any a 's
and any x , and for any x_1 's
 b , and any N and for any x_1 's

$$\mathbb{L}\left[\sum_{n=1}^N a_n x_n\right] + b = \sum_{n=1}^N a_n \mathbb{L}[x_n] + b$$

(Principle of Superposition).
For any two linear operators $\mathbb{L}[x]$ and

$$S(Y), \quad E[S(Y)] = S[E(Y)] \quad (3)$$

Expectation is linear. $x \rightarrow [S] \rightarrow [E] \rightarrow \phi$

pf for Theorem 2.12: $y \rightarrow [E] \rightarrow [S] \rightarrow \phi$

Let $Y = f(X) = aX + b$, Then, from

Theorem 2.10, m

$$E(Y) = \sum_{x \in S_X} f(x) P_X(x) = \sum_{x \in S_X} (ax + b) P_X(x)$$

$$= a \underbrace{\sum_{x \in S_X} x P_X(x)}_{M_X} + b \underbrace{\sum_{x \in S_X} P_X(x)}_{1} \quad \text{Q.E.D.}$$

If $W = g(R)$ where $g(x)$ is
nonlinear, then

$$E[W] \neq g(E[R])$$

If $W = g(R)$ where $g(x)$ is
linear, then

$$E[W] = g(E[R]).$$

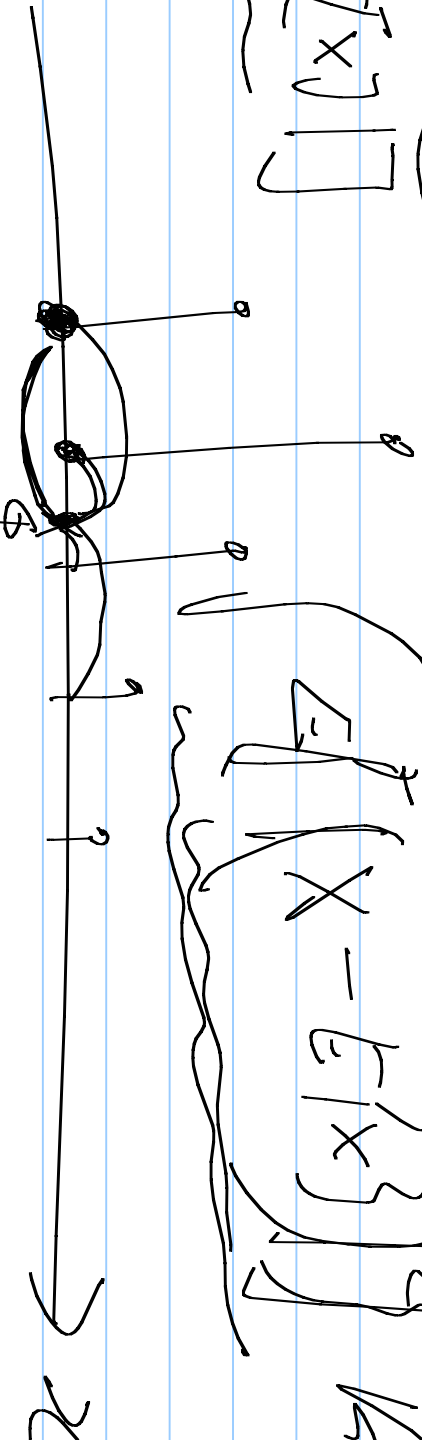
$$P_X(x)$$

$$E[X - E[X]] = 0$$

$$E[X - E[X]]$$

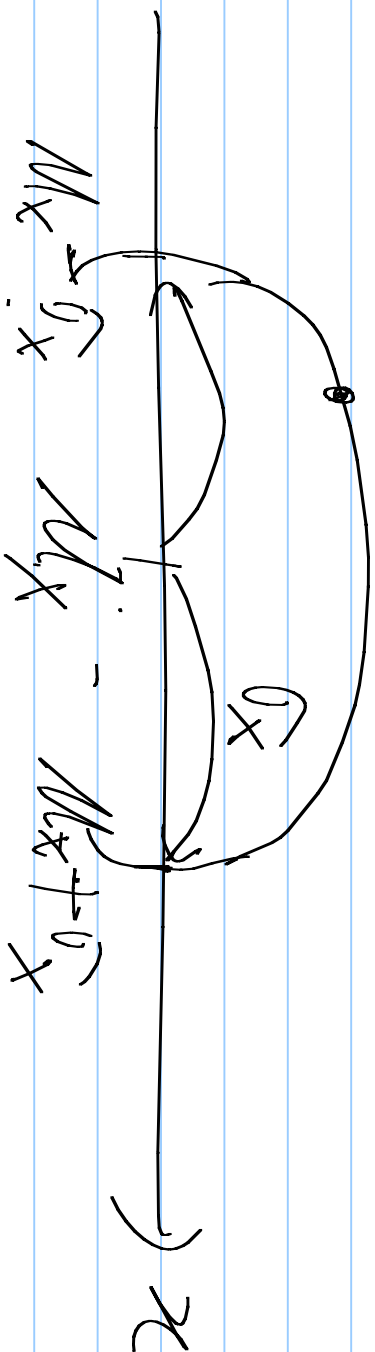
$$E[X - E[X]]$$

variance



$P_X(x)$ completely describes the probability model of X , and so does $F_X(x)$.

For some X ,



$$D[X \in (M_X - D_X, M_X + D_X)]$$

$\text{Var}[X]$ or D_X describes the dispersion of X relative to M_X .

$$\text{Var}[X] \equiv E[(X - \mu_X)^2]$$

$$\text{Law of } \sum_{x \in S_X} \underbrace{(x - \mu_X)^2}_{(x - \mu_X)^2 + \mu_X^2} = f_X(x)$$

Minimizations starts from

$E[X^2]$ is commonly called the mean square value of X .

All the moments $\{E[X^n] \mid n \in \mathbb{N}, 1, 2, \dots\}$
completely describe the probability
model of X .

$$\begin{aligned}\text{Var}[X] &= E[(X - \mu_X)^2] \geq 0 \\ &= E[X^2] - \mu_X^2 \geq 0\end{aligned}$$

$$\Rightarrow E[X^2] \geq \mu_X^2 \text{ for any}$$

random variable X .

Recall: $P[A|B]$ denotes the conditional probability of event A given that event B occurs, with $P[B] > 0$. That is, we define

$$P[A|B] = \frac{P[A \cap B]}{P[B]}.$$

If we let $A = \{X = x\}$ for a discrete random variable, then

$$P[A|B] = P[X = x|B] = P_{X|B}(x)$$

$$\underbrace{P_{X|B}(x)} \stackrel{\Delta}{=} P[X=x|B] = \frac{P[X=x, B]}{P[B]}$$

$P_{X|B}(a)$, $P_{X|B}(b)$
dummy argument

If $x \in B$,

$$P_{X|B}(x) = \frac{P[X=x, B]}{P[B]} =$$

$$= \frac{P[X=x]}{P[B]} = \frac{P_X(x)}{P[B]}$$

If $x \notin B$,

$$P_{X|B}(x) = \frac{0}{P[B]} = 0$$

$\theta \in \mathbb{R}$

Note that conditional expectation $E[X|B]$ satisfies the law of

total variance (i.e.,)

$$E[g(X)|B] = \sum_{x \in B} g(x) f_{X|B}(x), \text{ and}$$

is linear, i.e., for any constants a, b , $E[aX + b|B] = aE[X|B] + b$.

$\{X = x\}$ is an event representing

$\{S \in S \mid \underbrace{X(S) = x}_{\text{condition}}\}$
argument value

$\{X = x\} = \{S \mid X(S) = x\}$ S_A no countable

$\{A\} \equiv \#$ of students asleep (discrete)

$\{C\} \equiv \#$ of phone calls (C is countable)

$\{N\} \equiv \#$ of minutes (S_N is NOT countable)

Grade (Letter) \rightarrow Grade (Number)

A

4

B

3



C

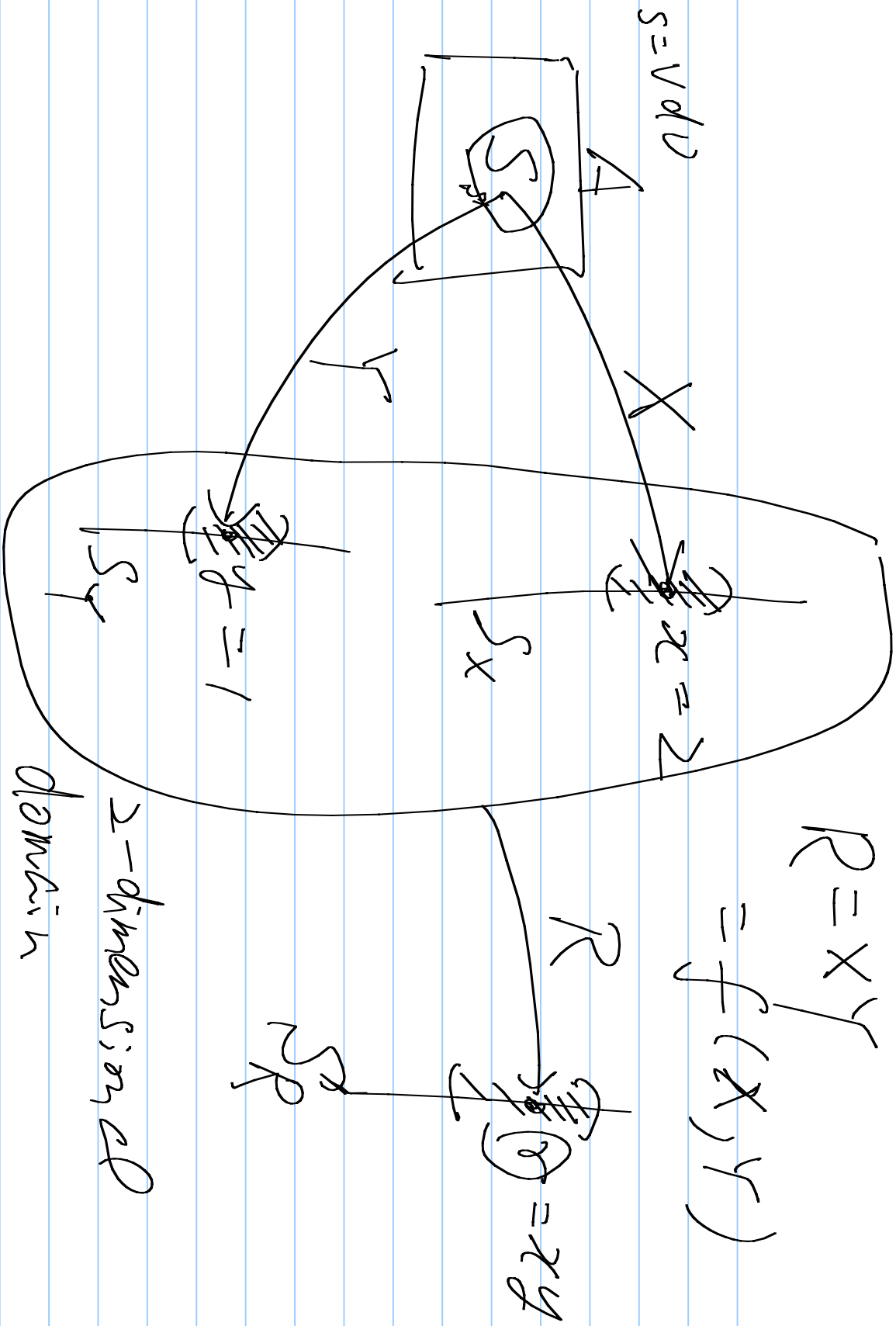
2

D

1

E

0



$$P_X(x) \equiv P[X=x]$$

$$= P\{s \mid X(s)=x\}$$

An event of outcomes
satisfying the
condition $X(s)=x$

Define event $A = \{X = x\}$

$$P[A] = P[X = x] = P_X(x)$$

Axiom 1: for any event A ,

$$P[A] \geq 0 \quad \text{and} \quad P_X(x) \geq 0$$

for any x

$$\text{Axiom 2: } P[S] = 1.$$

$$S = S_X \Rightarrow P[S] = P[X \in S_X] = \sum_{x \in S_X} P_X(x)$$

$$\sum_{x \in S} \mathbb{I} \left[\bigcup_{x \in S} \{X=x\} \right] = \sum_{x \in S} \mathbb{P}_x(x) = 1$$

Axiom 3 Axiom 2

c) For any $B \subset \mathcal{X}$, the probability that X is in the set B is

$$P[X \in B] = P[\bigcup_{x \in B} \{X=x\}]$$

$$\text{Axiom } \} \quad \overset{=}{\sum_{x \in B}} \underbrace{P[X=x]}_{P_X(x)}$$

$$= \sum_{x \in B} P_X(x) \quad \#$$

A statistic is a single number (nonrandom or random) that is derived from a probability model.

- * repeatable independent trials
- * sample space for single trial = $\{0, 1\}$

with $P[0] = 1-p$ and $P[1] = p$.

$$\underbrace{00\dots 01}_{1^0} P[X=x] = (1-p)^{x-1} p$$
$$\underbrace{1^p}_{x-1} \quad \underbrace{1^p}$$

Since $E[\cdot]$ is linear,

$$E[(X - \mu_x)^2] = E[(X^2 - 2\mu_x X + \mu_x^2)]$$

$$= E[X^2] - 2\mu_x \underbrace{E[X]}_{\mu_x} + \underbrace{E[\mu_x^2]}_{\mu_x^2 \text{ constant}}$$

$$= \underline{E[X^2] - \mu_x^2}$$

$$P[A] = \sum_{n=1}^m \underbrace{P[A \cap B_n^-]}$$

$$= \sum_{n=1}^m P[A|B_n^-] P[B_n^-]$$

$$X = \sum_{n=1}^m$$

