

CHAPTER TWO

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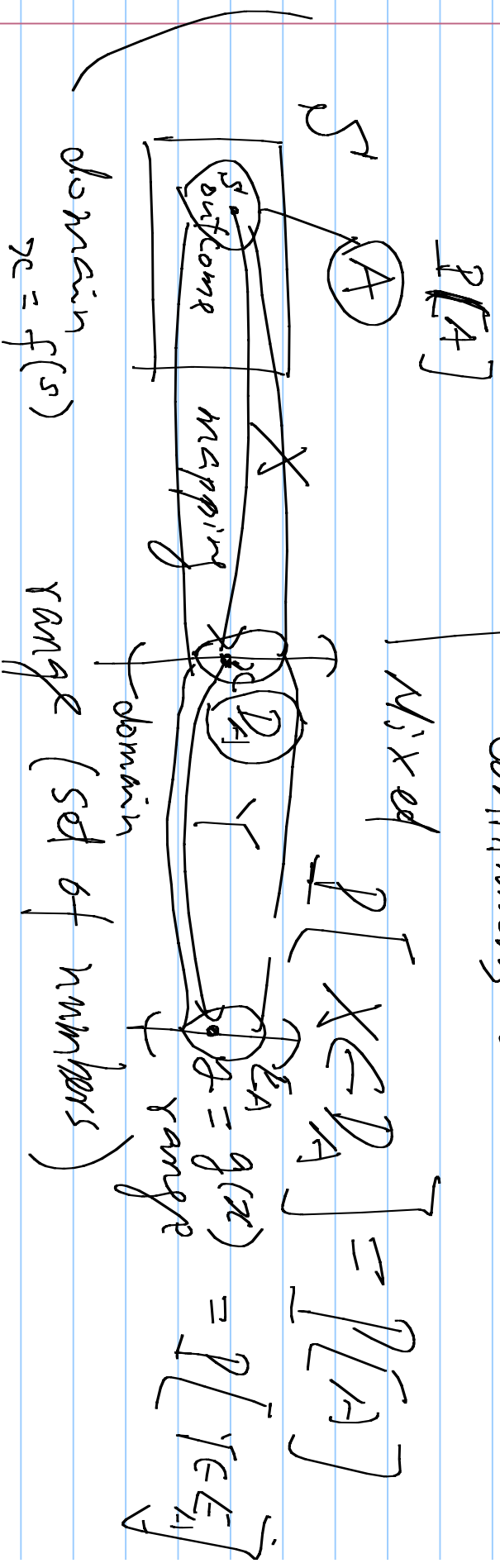
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Experiment: Procedure and Observation

Physical Model \Rightarrow Probability Model

Random Variable

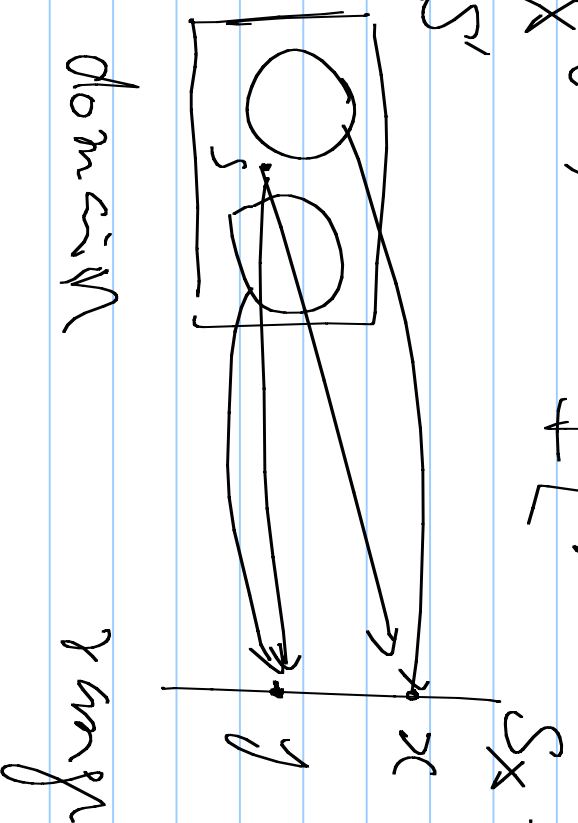
Discrete \checkmark
Continuous \checkmark



$$\boxed{P[X=x]} \equiv P[A] \text{ where}$$

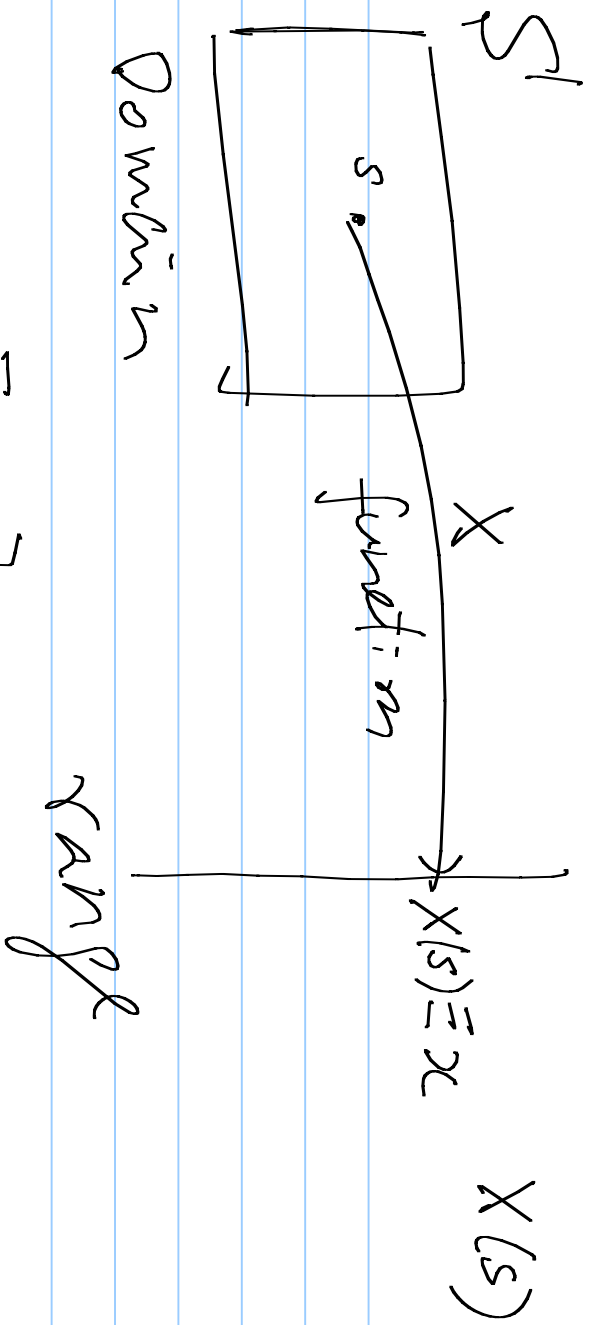
$$A = \{s \mid X(s) = x\}$$

$$P_X(x) \equiv P[X=x]$$



$$\{X=x\} \cap \{X=y\} = \emptyset$$

if $x \neq y$.



$$P_X(x) \stackrel{\text{def}}{=} P\left[\left. \begin{array}{l} X=x \\ \underbrace{P[S | X(s)=x]} \end{array} \right\} \right]$$

Recall: Consider ^{Repeatable} n experiment with identical and independent subexperiments.

Each subexperiment gives two possible outcomes: namely success, with probability p , and failure, with probability $1-p$.

Bernoulli (p) random variable represents the number of successes in one single trial

Geometric (p) random variable represents the number of trials until the first.

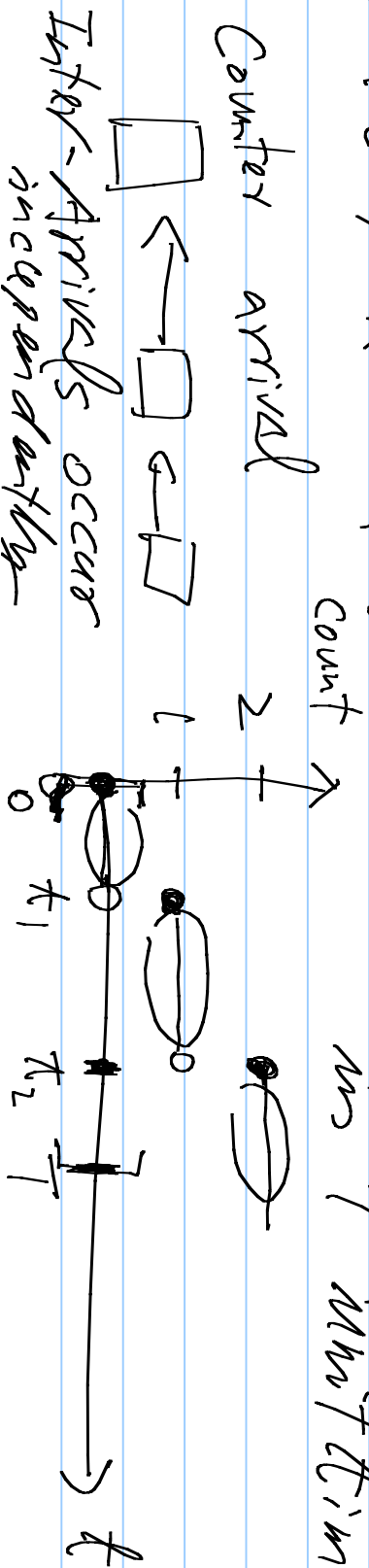
Success (inclusive)

Binomial (n, p) random variable represents

The number of successes in n trials.
Note: Binomial $(1, p) = \text{Bernoulli}(p)$
Pascal (k, p) random variable represents

the number of trials until k successes (inclusive). Note: Pascal $(1, p)$ is equivalent to Geometric (p) .

Poisson (λ) Random Variable = number of arrivals in T unit times

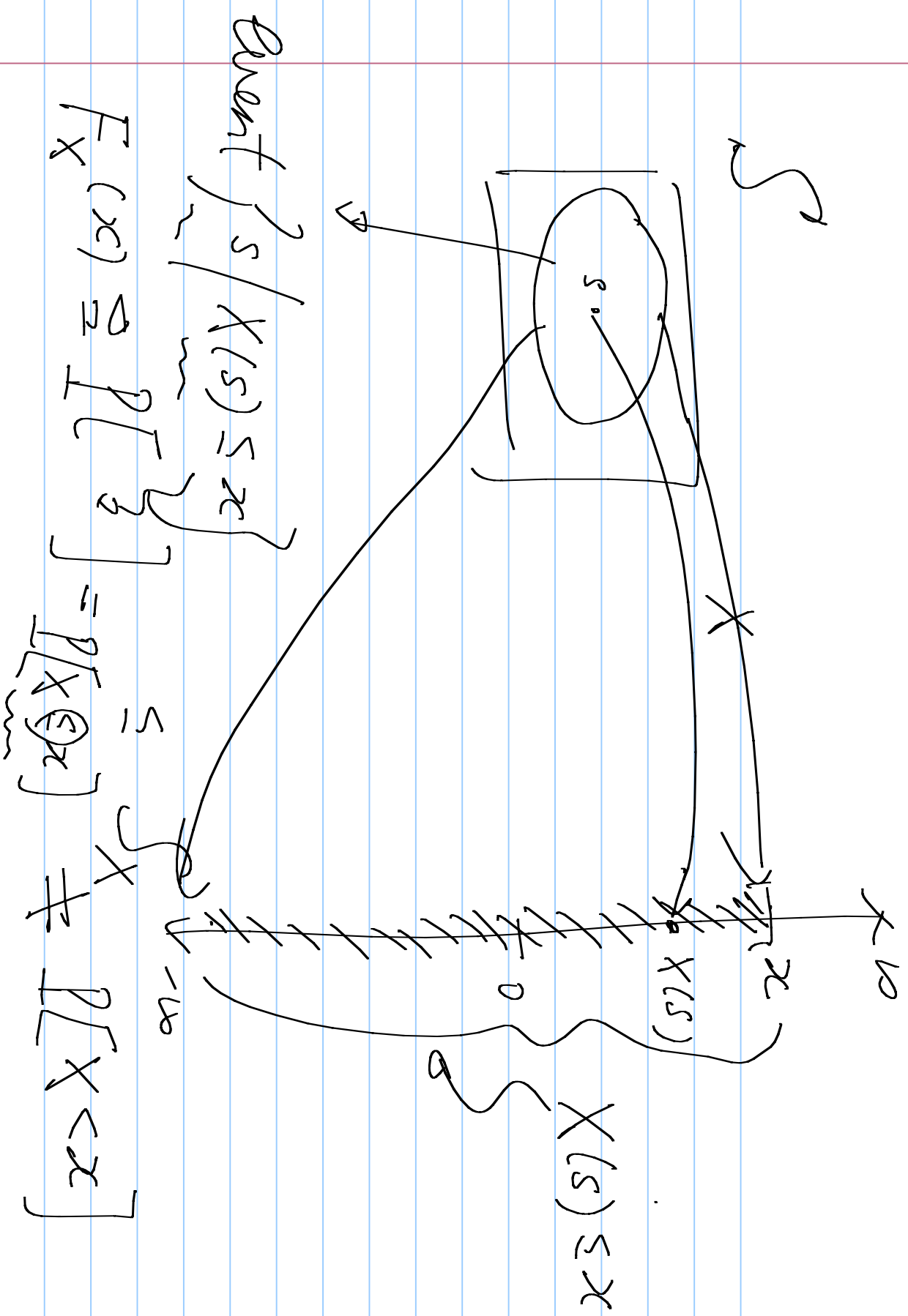


α denotes the average number of arrivals
in T unit times

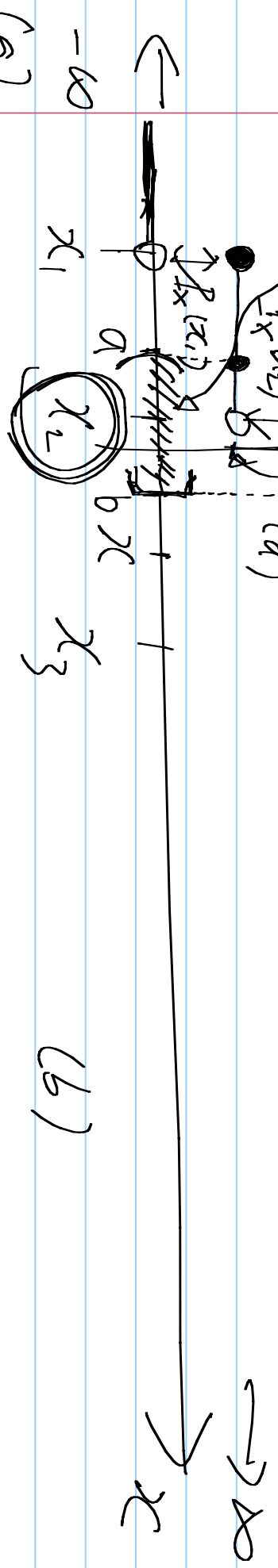
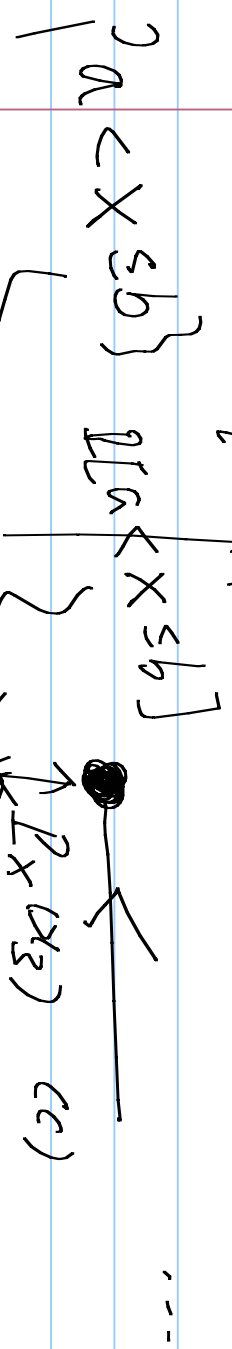
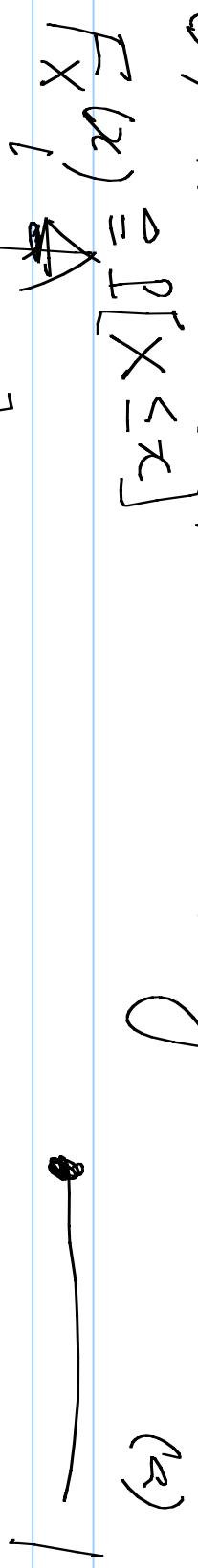
λ denotes the average number of arrivals
in unit time

$$P_k(R) = \frac{\alpha^k e^{-\alpha}}{k!}, \quad k=0, 1, \dots$$

$$\alpha = \lambda T$$

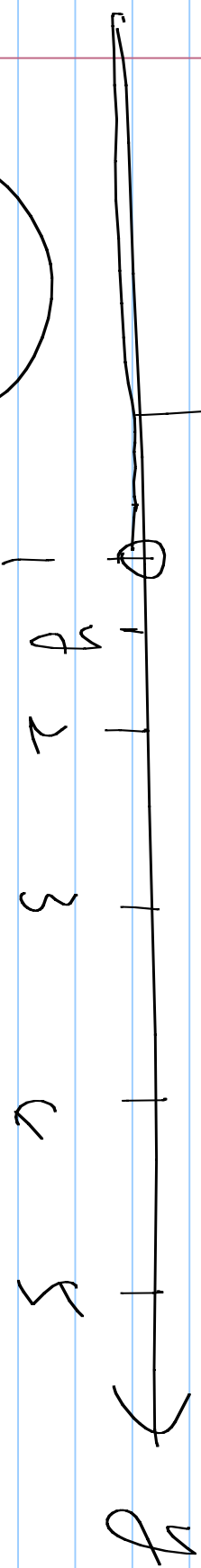
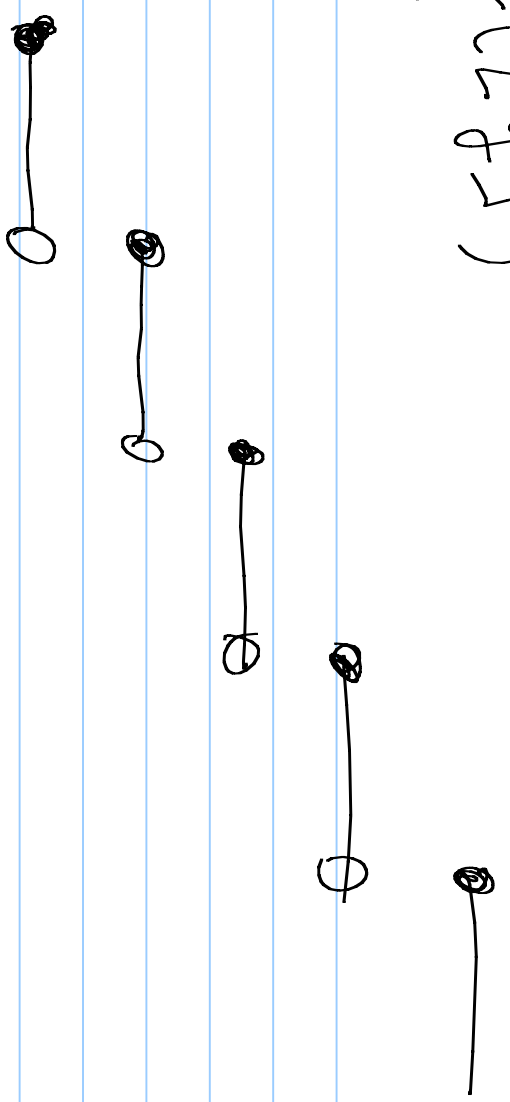


$F_X(x)$ must be nondecreasing.



(c) $F_X(x)$ is continuous from the right, but not necessarily from the left.

$$F_r(y) = F_r(\lfloor y \rfloor)$$



$\lfloor y \rfloor$ is the largest integer that is not larger than y .
 $\lfloor y \rfloor$ is integer part of y .

[y] \equiv the smallest integer that
is not smaller than y

• The average-value of a set of n experimental
outcomes is a statistic of the outcomes.

Three averages of interest:

mean (or expected value,
or expectation)

median

mode (unimodal, multimodal)

Consider an experiment that produces a rv X .
Perform n independent trials of this
experiment. Let x_i denote the i th
sample value.

$$M_n \equiv \text{sample average} = \frac{1}{n} \sum_{i=1}^n x_i$$
$$= \frac{1}{n} \sum_{x \in S_X} x N_x \quad (N_x \equiv \text{the number of sample values } x \text{ in } n \text{ trials})$$

$$M_n = \sum_{x \in S_X} x \frac{N_x}{n}$$

We can interpret $P[X=x]$ as $\lim_{n \rightarrow \infty} \frac{N_x}{n}$
(relative frequency interpretation)

$$\begin{aligned}
 \lim_{n \rightarrow \infty} M_n &= \lim_{n \rightarrow \infty} \sum_{x \in S_X} x \frac{d(x)}{n} \\
 &= \sum_{x \in S_X} x \underbrace{\lim_{n \rightarrow \infty} \frac{1}{n}}_{P[X=x]} \\
 &= \sum_{x \in S_X} x P[X=x] \\
 &= E[X]
 \end{aligned}$$

$$\sum_{x=0}^{\infty} r^x = \frac{1}{1-r} \quad \text{for } 0 < r < 1$$

$$\sum_{x=1}^{\infty} x r^x = \sum_{x=0}^{\infty} x r^x = \sum_{x=0}^{\infty} r \frac{d}{dr} r^x$$

$$= r \frac{d}{dr} \sum_{x=0}^{\infty} r^x = r \frac{d}{dr} (1-r)^{-1} = \frac{1}{1-r} + \frac{(1-r)^{-2}}{-2}$$

$$\Rightarrow \frac{r}{(1-r)^2} = \frac{r}{(1-r)^2} = \frac{1}{1-r}$$

$$\lim_{n \rightarrow \infty} P_k(R) = \lim_{n \rightarrow \infty} \frac{n(n-1) \dots (n-k+1)}{n^k}$$

1/

= 1

$$\cdot \left(\frac{\alpha^k}{R^k} \right) \cdot \lim_{n \rightarrow \infty} \left(1 - \frac{\alpha}{n} \right)^{n-k}$$

$\frac{\alpha^k}{R^k}$

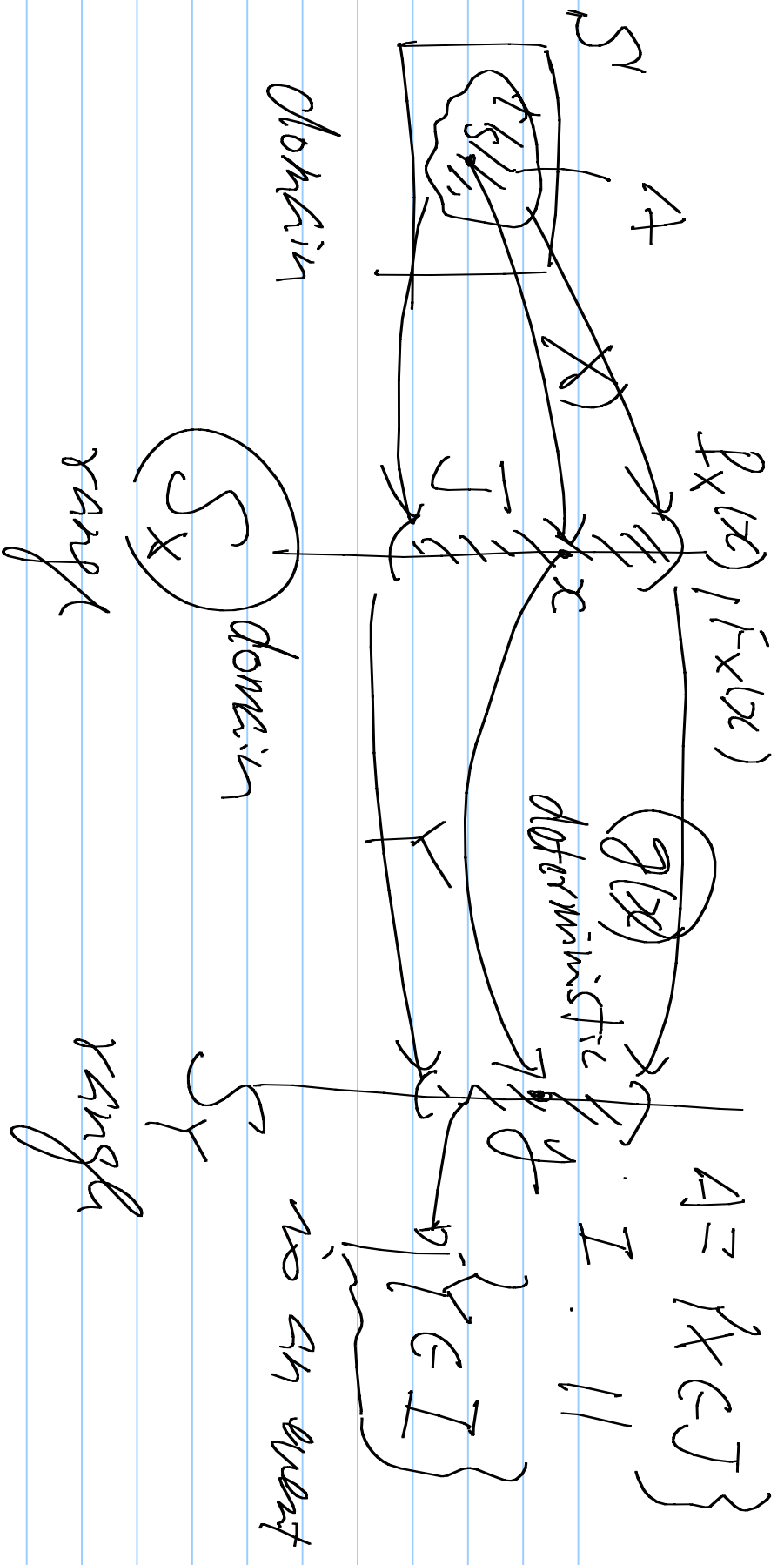
$$= \lim_{n \rightarrow \infty} \left(1 - \frac{\alpha}{n} \right)^n$$

$$\lim_{n \rightarrow \infty} \left(1 - \left(\frac{\alpha}{n} \right) \right)^n = 1$$

$$= \lim_{n \rightarrow \infty} \left(1 - \frac{\alpha}{n} \right)^n = \lim_{n \rightarrow \infty} e^{n \ln \left(1 - \frac{\alpha}{n} \right)}$$

$$\frac{0}{0}$$

$$\begin{aligned}
 L'_{\text{Hopital}} &= \lim_{h \rightarrow 0} \frac{h \ln(1 - \frac{\alpha}{h})}{\frac{\alpha}{h}} = \lim_{h \rightarrow 0} \frac{\ln(1 - \frac{\alpha}{h})}{1/h} \\
 &= \lim_{h \rightarrow 0} \frac{\frac{1 - \frac{\alpha}{h}}{-h^{-2}}}{-1/h} = \lim_{h \rightarrow 0} \frac{-\alpha}{1 - \frac{\alpha}{h}} = -\alpha
 \end{aligned}$$



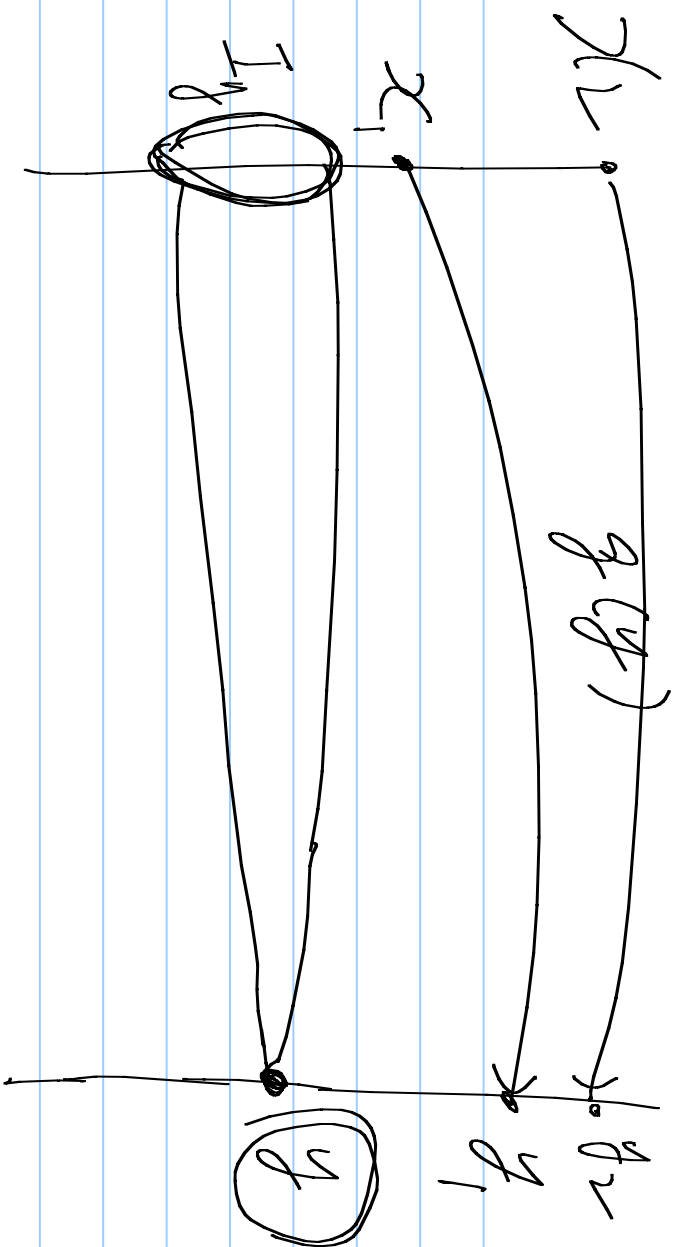
X is completely statistically characterized

if $P_X(x)$ or $F_X(x)$ is known.
 $A = \{s | X(s) \in J\} = \{s | Y(s) \in I\}$

Y is called the derived random

variable.

Denote $Y = g(X)$ ✓

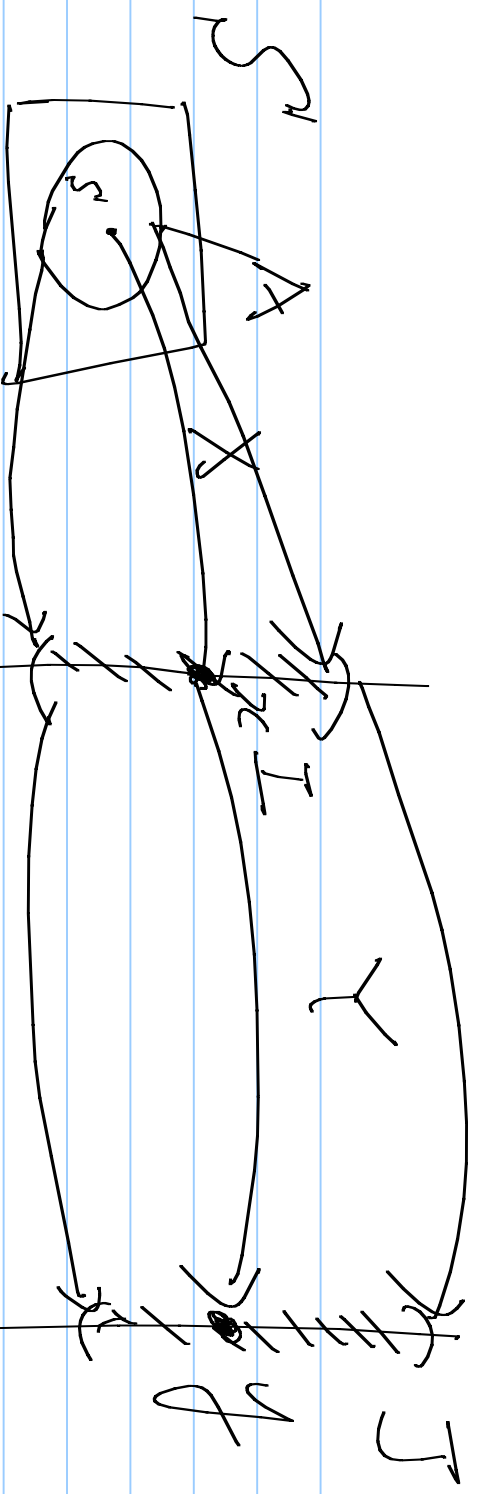


S_X

S_Y

- ① When $g(x_1) \neq g(x_2)$ for $x_1 \neq x_2$ (i.e., $g(x)$ is one-to-one), $P_Y(y) \stackrel{!}{=} P_X(x)$ if $y = g(x)$.
- ② When $y = g(x)$ for all $x \in I_y \stackrel{!}{=} \{x \mid g(x) = y\}$.

$$P_Y(y) = \sum_{x \in I_y} \underbrace{P_X(x)}$$



domain

$$A = \left\{ s \mid \left. \begin{array}{l} X(s) \in I \\ S_X \end{array} \right\} \right.$$

$$= \left\{ X \in I \right\}$$

range

$$= \left\{ Y(X(s)) \in J \right\}$$

domain

$$= \left\{ Y \in J \right\}$$

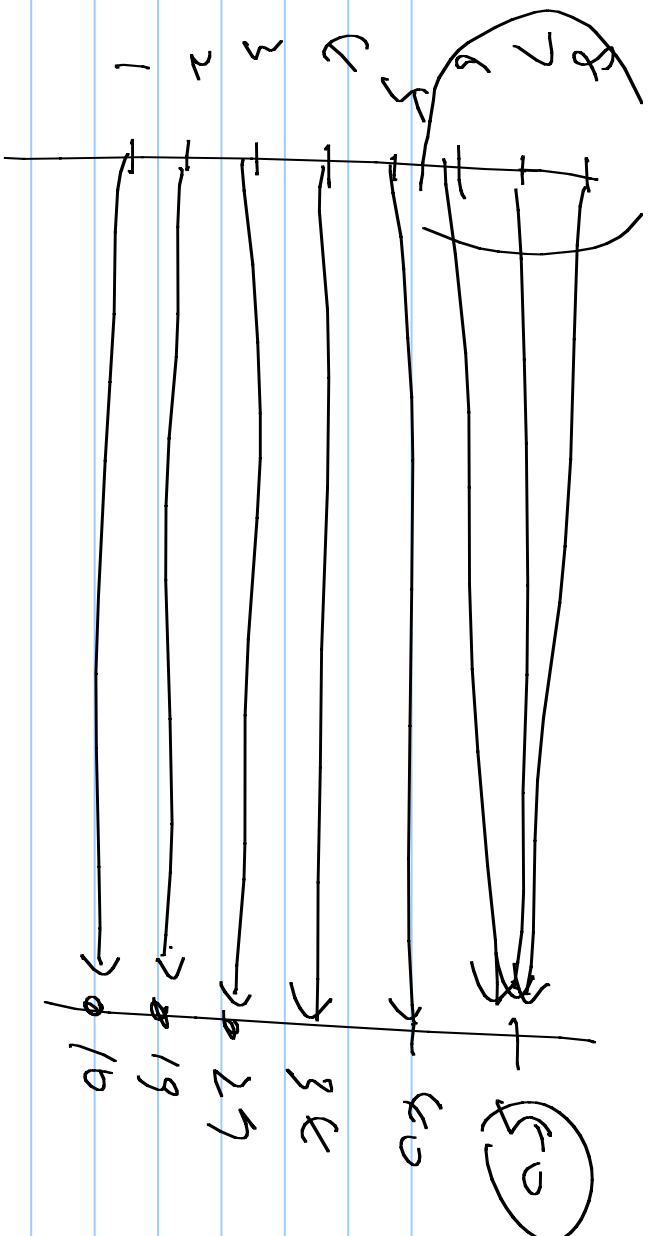
range

$$(Y = g(X))$$

When $\{I = y\}$, $A = \{x \mid g(x) = y\} = \{s \mid g(\underbrace{x(s)}) = y\}$ and

$$\underbrace{P[Y = y]} = \underbrace{P[X = y]} = \underbrace{P[A]} = \sum_{x: g(x) = y} P_X(x)$$

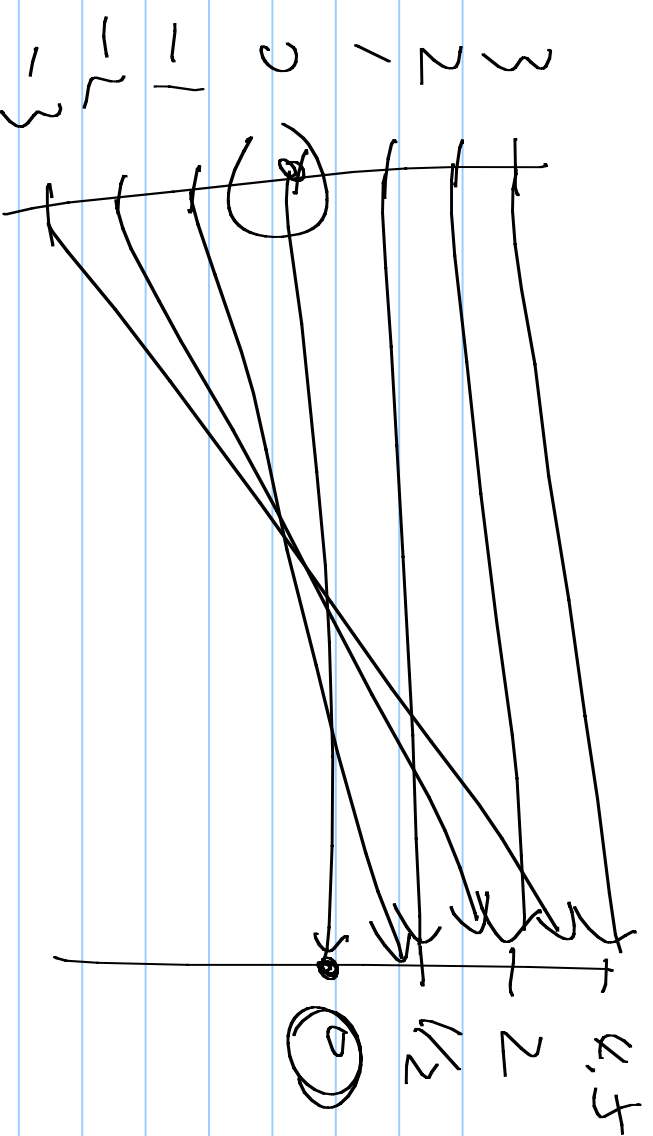
$$P_Y(y)$$



S_x S_y

Many - to - one mapping

Given $P_x(x)$ and $y = g(x)$, this section deals with the approach of finding $P_y(y)$.



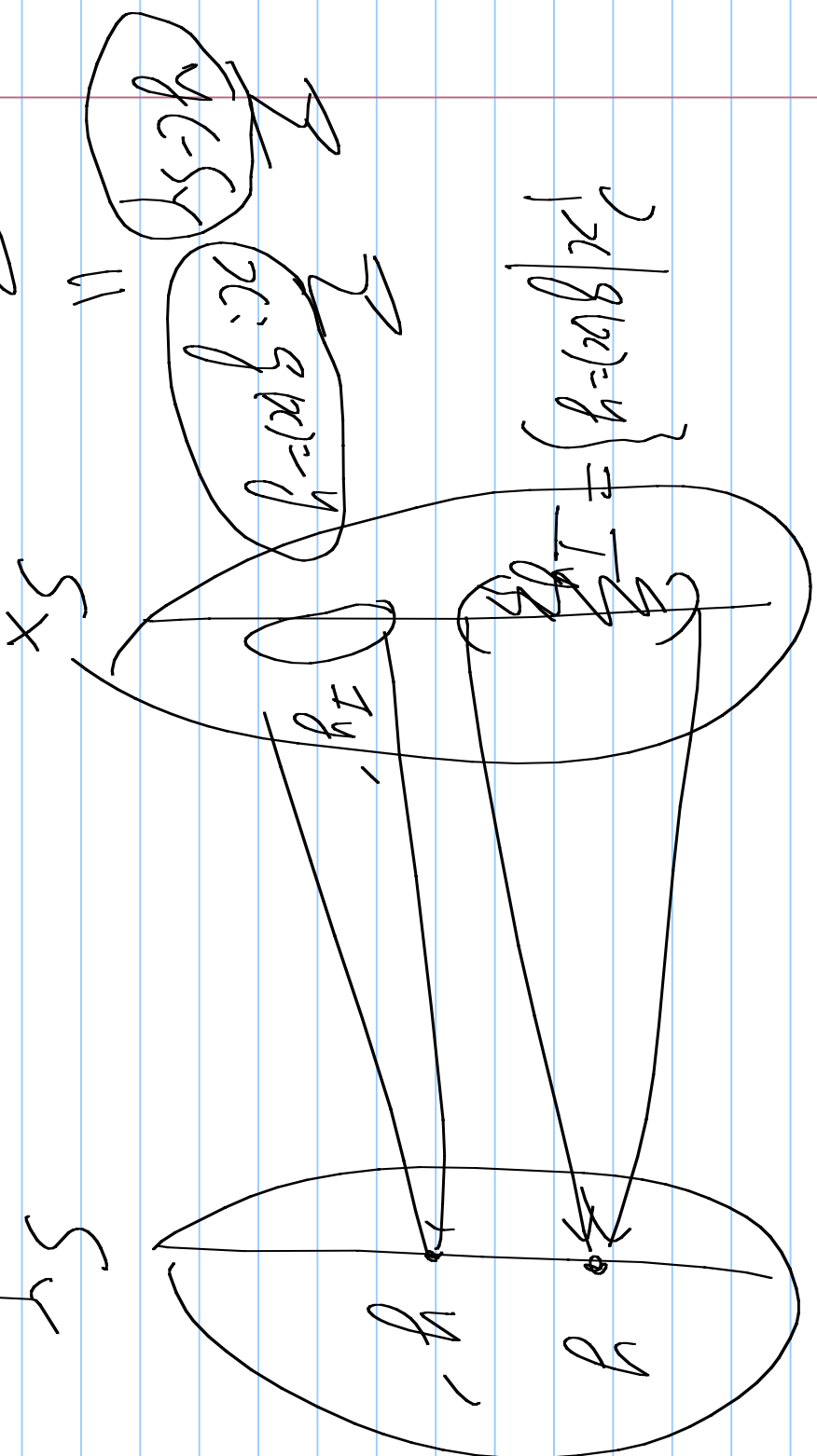
S_X

S_Y

$$\left. \begin{aligned} \{Y=0\} &= \{X=0\}, \{Y=1\} = \{X \in \{-1, 1\}\} \\ \{Y=2\} &= \{X \in \{-2, 2\}\}, \{Y=4.5\} = \{X \in \{-3, 3\}\} \end{aligned} \right\}$$

Law of Unconscious Statistician

$$P\{X | g(X) = y\} = \frac{f_X(x)}{f_Y(y)}$$



$$\sum_{x \in S_X}$$

* An operator $\mathbb{T}[x]$ is called

Linear (affine) iff for any a_n 's
and any x_n and for any x_n 's

$$\mathbb{T}\left[\sum_{n=1}^N a_n x_n\right] = \sum_{n=1}^N a_n \mathbb{T}[x_n]$$

(Principle of Superposition).

For any two linear operators $\mathbb{T}[x]$ and

$$S[Y], \quad E[S[Y]] = S[E[Y]]$$

Expectation is linear.

pf for Theorem 2.12:

Let $Y = f(X) = aX + b$, then, from

Theorem 2.10,

$$E[Y] = \sum_{x \in S_X} f(x) P_X(x) = \sum_{x \in S_X} (ax + b) P_X(x)$$

$$= a \underbrace{\sum_{x \in S_X} x P_X(x)}_{M_X} + b \underbrace{\sum_{x \in S_X} P_X(x)}_{\substack{1 \\ \text{Q.E.D.}}}$$

If $W = g(R)$ where $g(x)$ is
nonlinear, then

$$E[W] \neq g(E[R])$$

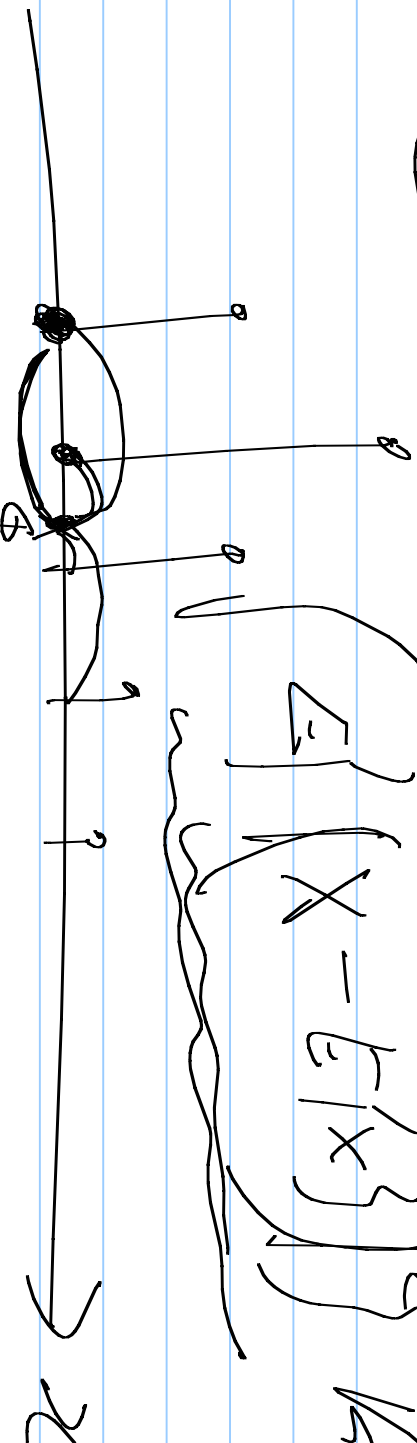
If $W = g(R)$ where $g(x)$ is
linear, then

$$E[W] = g(E[R]).$$

$$P_X(x)$$

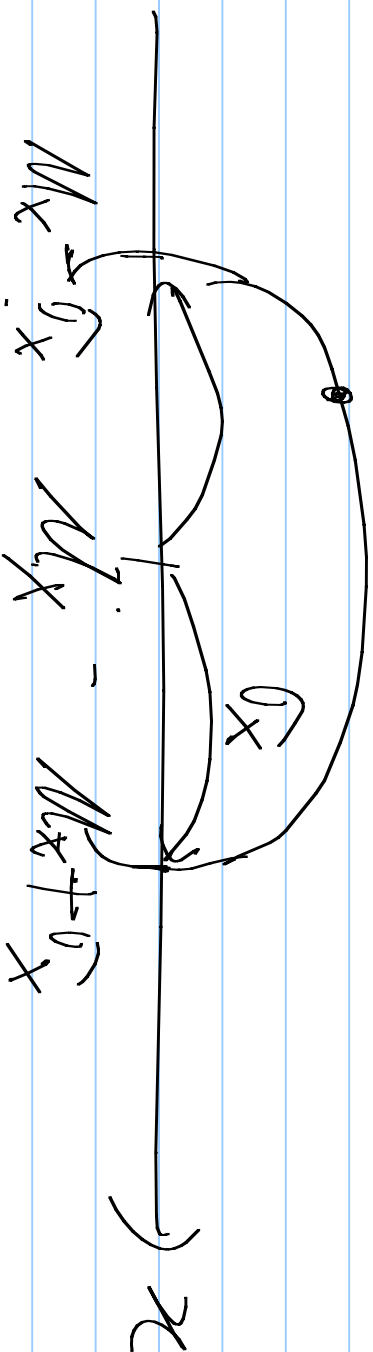
$$E\{X - E\{X}\} = 0$$

$$E\{X - E\{X\}\}^2 = \text{variance}$$



$P_X(x)$ completely describes the probability model of X , and so does $E\{X\}$.

For some X ,



$$P[X \in (M_X - \sigma_X, M_X + \sigma_X)]$$

$\text{Var}[X]$ or σ_X^2 describes the dispersion of X relative to M_X .

$$\text{Var}[X] \equiv E[(X - \mu_X)^2]$$

$$\text{Law of } \sum_{x \in S_X} \underbrace{(x - \mu_X)^2}_{(x - \mu_X)^2 + \mu_X^2 - 2\mu_X x} = f_X(x)$$

Non constant
starts from \equiv

$E[X^2]$ is commonly called the mean square value of X .

All the moments $\{E[X^n] | n \in \mathbb{N}, 1, 2, \dots\}$
completely describe the probability
model of X .

$$\begin{aligned}\text{Var}[X] &= E[(X - M_X)^2] \geq 0 \\ &= E[X^2] - M_X^2 \geq 0\end{aligned}$$

$$\Rightarrow E[X^2] \geq M_X^2 \text{ for any}$$

random variable X .

Recall: $P[A|B]$ denotes the conditional probability of event A given that event B occurs, with $P[B] > 0$. That is, we define

$$P[A|B] = \frac{P[A \cap B]}{P[B]}.$$

If we let $A = \{X = x\}$ for a discrete random variable, then

$$P[A|B] = P[X = x|B] = P_{X|B}(x)$$

$$P_{X|B}(x) \equiv P[X=x|B] = \frac{P[X=x, B]}{P[B]}$$

$P_{X|B}(a)$, $P_{X|B}(b)$
dummy argument

If $x \in B$,

$$P_{X|B}(x) = \frac{P[X=x, B]}{P[B]} =$$

$$= \frac{P[X=x]}{P[B]} = \frac{P_X(x)}{P[B]}$$

If $x \notin B$,

$$P_{X|B}(x) = \frac{0}{P[B]} = 0$$

$\theta \in \mathbb{R}$

Note that conditional expectation $E[X|B]$ satisfies the law of

total expectations, i.e.,

$$E[g(X)|B] = \sum_{x \in B} g(x) f_{X|B}(x), \text{ and}$$

is linear, i.e., for any constants a, b , $E[aX + b|B] = aE[X|B] + b$.

